

# USING RANDOM WALKS TO DETECT AMENABILITY IN F

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Murray Elder, Andrew Rechnitzer, Buks van Rensburg, Cameron Rogers  
Odense, August 2016

[4] Burillo, Cleary and Wiest 2007

The authors randomly choose words and reduce them to a normal form to test if they represent the identity element. From this they estimate the proportion of words of length  $n$  equal to the identity, as a way to compute the asymptotic growth rate of the cogrowth function.

[1] Arzhantseva, Guba, Lustig, and Préaux 2008

The authors study the *density* or least upper bound for the average vertex degree of any finite subgraph of the Cayley graph; an  $m$ -generated group is amenable if and only if the density of the corresponding Cayley graph is  $2m$ . They use a computer program to find a finite subset in  $F$  with density 2.89577. (To be amenable one would need to find sets whose density  $\rightarrow 4$ ).

[6] Elder, Rechnitzer and Wong 2012

Lower bounds on the cogrowth rates of various groups are obtained by computing the dominant eigenvalue of the adjacency matrix of truncated Cayley graphs. These bounds are extrapolated to estimate the cogrowth rate. As a byproduct the first 22 coefficients of the cogrowth series are computed exactly.

[10] Haagerup, Haagerup, and Ramirez-Solano 2015

Lower bounds on the norms of the Markov operator derived from Kesten's condition are obtained for  $F$  using  $C^*$ -algebraic methods. Coefficients of the cogrowth series are computed exactly to 48 terms.

[5] Elder, Rechnitzer and van Rensburg 2015

The *Metropolis Monte Carlo* method from statistical mechanics is adapted to estimate the asymptotic growth rate of the cogrowth function by running random walks on the set of all trivial words in a group. The results obtained for Thompson's group  $F$  suggest it to be non-amenable.

Justin Moore [12] (2013) has shown that if  $F$  were amenable then its Følner function would increase faster than a tower of  $n - 1$  twos,

$$2^{2^{\dots}}$$

This has been proposed as an obstruction to all computational methods for approximating amenability; a computationally infeasibly large portion of the Cayley graph must be considered before sets with small boundaries can be found.

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However, in all but one of the experimental algorithms listed above computing Følner sets was not the principle aim. Exactly how the growth of the Følner function controls the convergence properties of the respective limits in the Grigorchuk-Cohen, Kesten, Reiter, ... characterisations is not clear.

## COGROWTH

$G =$  a group;  $S = S^{-1}$  = a finite generating set for  $G$ .

$d_n = \#$  words of length  $n$  in  $S^*$  equal to  $e$ .

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Moves are accepted/rejected with a carefully chosen probability (depending on the relative change in length of the current state, and parameters  $\alpha, \beta$ ).

Let  $w$  be the current state (a reduced trivial word).

- If  $w'$  was obtained from  $w$  via a conjugation it is accepted as the new state with probability

$$\min \left\{ 1, \left( \frac{|w'| + 1}{|w| + 1} \right)^{1+\alpha} \beta^{|w'| - |w|} \right\}.$$

- If  $w'$  was obtained from  $w$  via an insertion it is accepted as the new state with probability

$$\min \left\{ 1, \left( \frac{|w'| + 1}{|w| + 1} \right)^\alpha \beta^{|w'| - |w|} \right\}.$$

## Example

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Note:

- random walk on trivial words, not the Cayley graph
- just need presentation, no efficient normal form or even solvable word problem

If  $\Pr(u \rightarrow v)$  is the probability of moving from  $u$  to  $v$  in one step, a distribution  $\pi$  is *stationary* for the walk if  $\pi(u) = \sum_v \Pr(v \rightarrow u)\pi(v)$ .



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**Theorem (E, Rechnitzer, van Rensburg [5])**

$$\pi(w) = \frac{(|w| + 1)^{1+\alpha} \beta^{|w|}}{Z}$$

where  $Z$  is a normalising constant  
is the unique stationary distribution for the algorithm.

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*i.e.* the probability that the algorithm reaches state  $w$  after  $N$  steps converges to  $\pi(w)$ .

Using this we can test for amenability by estimating the location of the asymptotic growth rate for  $c_n$ , as follows.

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Expected length of a state in the walk is

$$\begin{aligned} E(|w|) &= \sum |w| \pi(w) = \sum |w| \frac{(|w| + 1)^{1+\alpha} \beta^{|w|}}{Z} \\ &= \sum_{n=0}^{\infty} \frac{n(n+1)^{1+\alpha} \beta^n}{Z} c_n \end{aligned}$$

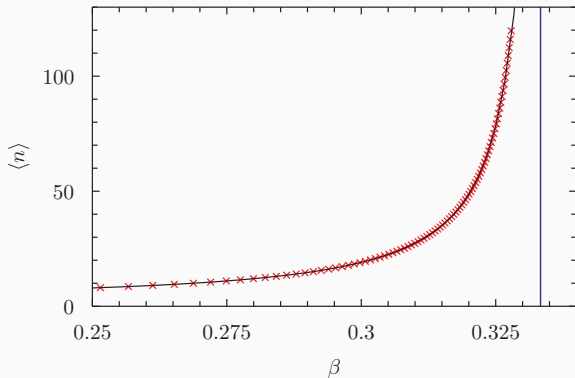
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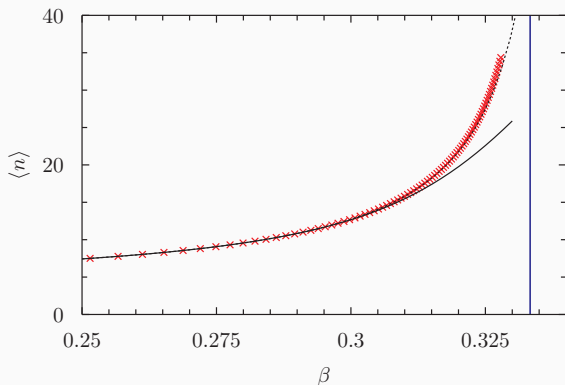
As  $\beta \rightarrow$  reciprocal of the cogrowth rate, the mean length changes from finite to infinite.

# FREE ABELIAN GROUP



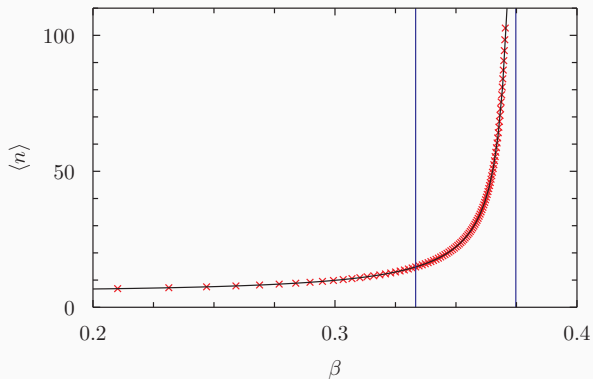
Mean length in  $\mathbb{Z}^2 = \langle a, b \mid bab^{-1}a^{-1} \rangle$  with  $\alpha = 1$ .

Reciprocal of  $|S| - 1$  is  $1/3$ .



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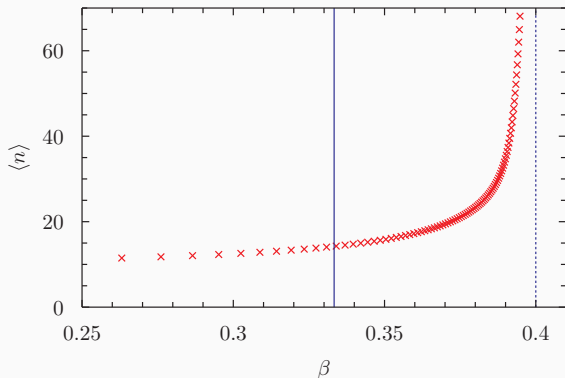


Mean length in  $BS(2, 2) = \langle a, b \mid ba^2b^{-1}a^{-2} \rangle$  with  $\alpha = 1$ .

Reciprocal of  $|S| - 1$  is  $1/3$ .



# THOMPSON'S GROUP F



Mean length in  $F = \langle a, b \mid [ab^{-1}, a^{-1}ba], [ab^{-1}, a^{-2}ba^2] \rangle$  with  $\alpha = 2$ .

Reciprocal of  $|S| - 1$  is  $1/3$ .

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Main objection to ERR:

### Theorem (Moore [12])

*If Thompson's group  $F$  is amenable it has a Følner function which grows like*

$$2^{2^{\dots}}$$

## Theorem (Følner)

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Vershik defined the following function as way of quantifying the rate of convergence of this limit.

$$\mathcal{F}(n) = \min \left\{ |F| : \frac{|\partial F|}{|F|} < \frac{1}{n} \right\}.$$

## Cameron

- ran his code on more groups, and for larger  $\alpha$  values
- collected extra data in addition to the mean length
- looked for a quantitative connection between the rate of convergence of ERR walks to their theoretical stationary distribution, and the Følner function.

Recall

- If  $w'$  was obtained from  $w$  via an insertion it is accepted as the new state with probability

$$\min \left\{ 1, \left( \frac{|w'| + 1}{|w| + 1} \right)^\alpha \beta^{|w'| - |w|} \right\}.$$



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By varying  $\alpha$  we can push the walk out to visit longer words.

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For large  $n$  the Metropolis rule will reject insertions of the long relator often. So you are essentially walking on the *wrong group*.

Recall: insert accepted with probability

$$\min \left\{ 1, \left( \frac{|w'| + 1}{|w| + 1} \right)^\alpha \beta^{|w'| - |w|} \right\}.$$

## WALKING ON THE WRONG GROUP

$n$	#steps	#short accepted	#long accepted
2	$3.6 \times 10^8$	4420185	5579815
3	$6.1 \times 10^8$	6323376	3676624
4	$9.0 \times 10^8$	8016495	1983505
5	$1.2 \times 10^9$	9088706	911294
6	$1.4 \times 10^9$	9621402	378598
7	$1.5 \times 10^9$	9850251	149749
8	$1.7 \times 10^9$	9943619	56381
9	$1.8 \times 10^9$	9977803	22197
10	$1.9 \times 10^9$	9991680	8320
11	$2.1 \times 10^9$	9997122	2878
12	$2.2 \times 10^9$	9998720	1280
13	$2.2 \times 10^9$	9999585	415
14	$2.3 \times 10^9$	9999938	62
15	$2.4 \times 10^9$	10000000	0
16	$2.6 \times 10^9$	10000000	0
17	$2.7 \times 10^9$	10000000	0

$$\langle a, b \mid aba = bab, a^n = b^{n+1} \rangle$$

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The long relator issue is easily detected – keep count of number of insertions of different relators. Rules out feasibility of method on infinitely presented groups.

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So no problem here.



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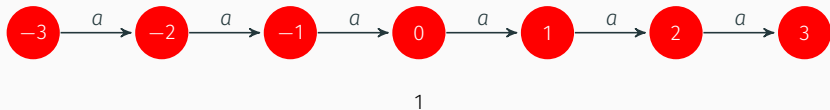
*If  $S = S^{-1}$  is a finite generating set, put  $\mu(x) = \frac{1}{|S|}$  for each  $x \in S$ .*

Then  $\mu$  induces a random walk on  $G$ :

- start at  $X_0 = e$ ;
- $X_{n+1} = X_n g$  where  $g \in G$  is chosen with probability  $\mu(g)$ .

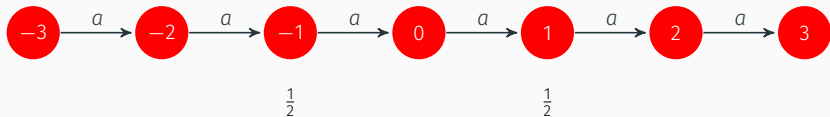
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$G = \mathbb{Z} = \langle a \rangle$ . Put  $\mu(a^{\pm 1}) = \frac{1}{2}$ :



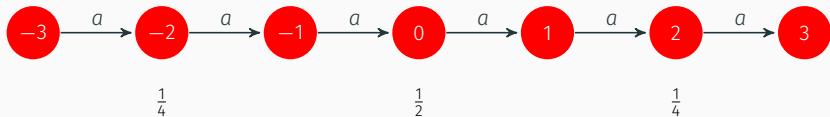
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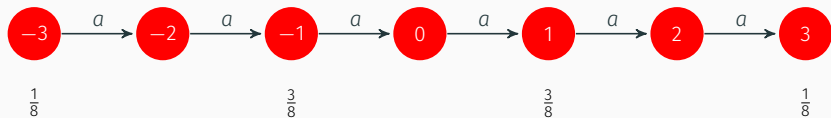
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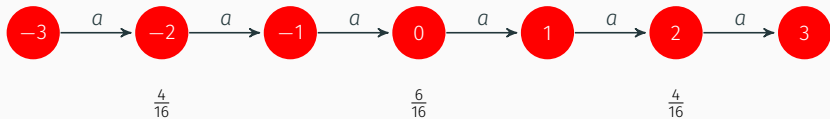
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Recall  $\sigma * \tau(g) = \sum_{h \in G} \sigma(h)\tau(h^{-1}g)$  is the *convolution* of two measures, so

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Note  $\mu_n(e) = \frac{d_n}{|S|^n}$  when  $\mu$  is the uniform measure on  $S$ .

**Lemma**

If  $\mu$  is symmetric with  $\text{supp}(\mu)$  generating  $G$ , then

- $\mu_{2n}$  is maximised at  $e$
- $\mu_{2n}(e)$  is non-increasing.

**Theorem (Avez [2])**

$G$  is amenable if and only if  $\frac{\mu_{2n}(x)}{\mu_{2n}(e)} \rightarrow 1$  for all  $x \in G$ .

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**Definition (E, Rogers [8])**

Let  $p \in (0, 1)$ . Define

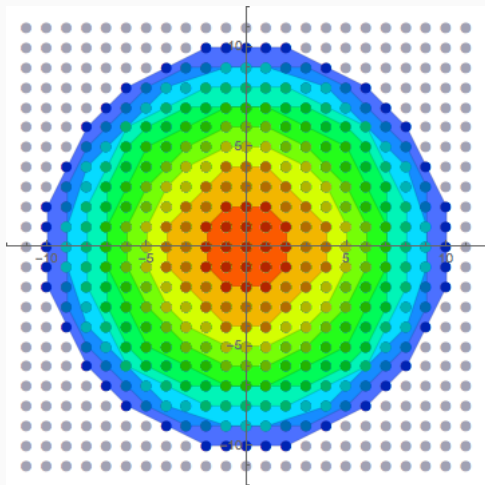
$$A_{\mu,n,p} = \left\{ g \in G \mid \frac{\mu_{2n}(g)}{\mu_{2n}(e)} > p \right\}$$

## EXAMPLE

$A_{\mu, n, \frac{1}{2}}$  for  $\mathbb{Z}^2 = \langle a, b \mid ab = ba \rangle$  with  $\mu(a^{\pm 1}) = \mu(b^{\pm 1}) = \mu(e) = \frac{1}{5}$ :

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Kamainovich and Vershik had shown a quantitative link the other direction:

### Theorem (Kaimanovich, Vershik [11])

*If  $F$  is a finite subset of  $G$  for which  $|F \cup Fs \setminus F \cap Fs| < \epsilon$  for every  $s \in S$  then for every  $p < 1$*

$$\mu_{2n}(e) \geq (1 - p)^n \frac{1 - 2\epsilon/p^2}{|F|}.$$

## MORE ON THESE SETS

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**Conjecture**

For any  $G$ , the set  $A_{\mu}$  is the largest normal amenable subgroup of  $G$  (the amenable radical).

Proofs use an alternative characterisation of  $A_\mu$  using:

### Theorem (Reiter [13])

$G$  is amenable iff for any finite subset  $K$  there is a sequence of unit vectors  $f_n \in L^2(G)$  such that

$$\lim_{n \rightarrow \infty} \|k \cdot f_n - f_n\|_2 = 0$$

for every  $k \in K$ .

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$G$  is amenable iff for any finite subset  $K$  there is a sequence of unit vectors  $f_n \in L^2(G)$  such that

$$\lim_{n \rightarrow \infty} \|k \cdot f_n - f_n\|_2 = 0$$

for every  $k \in K$ .

Where do we find with such a sequence of unit  $L^2$  functions? How about  $\mu_n(e)$ ?

### Proposition

Let  $\xi_n = \frac{\mu_n}{\|\mu_n\|_2}$ , then  $A_\mu = \left\{ g \in G \mid \lim_{n \rightarrow \infty} \|g \cdot \xi_n - \xi_n\|_2 = 0 \right\}$ .

## QUANTIFYING RATES OF CONVERGENCE

For amenable groups the probability of return is sub-exponential, and thus identifies the principle sub-dominant term in the asymptotics of the cogrowth function.

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The following definition quantifies this sub-dominant behaviour in a way analogous to the Følner function.

### Definition (E, Rogers [7])

Let  $D = \limsup d_n^{1/n}$  ( $= |S|$  when  $G$  is amenable). Define

$$\mathcal{R}(n) = \min \left\{ k : \frac{d_{2k+2}}{d_{2k}} > D^2 - \frac{1}{n} \right\}$$

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### Example

*For the trivial group with some finite symmetric generating set  $S$  we have  $d_n = |S|^n$  so  $\frac{d_{2k+2}}{d_{2k}} = |S|^2$ , so  $\mathcal{R}(n) = 0$ .*

Recall  $d_n = \mu_n(e)|S|^n$  where  $\mu$  is the uniform measure on  $S$ , so we can rephrase the set as

$$\mathcal{R}(n) = \min \left\{ k : \frac{\mu_{2k+2}(e)}{\mu_{2k}(e)} > 1 - \frac{1}{|S|^{2n}} \right\}.$$



## QUANTIFYING RATE OF CONVERGENCE OF COGROWTH

Return probabilities have been computed for various groups, and from these we can derive asymptotic formulae for  $\mathcal{R}(n)$ :

$G$	$\mathcal{F}(n)$	$\mu_n(e)$	$\mathcal{R}(n)$
trivial, $C_2$	$\asymp$ constant	$\asymp$ constant	$\asymp$ constant
finite	$\asymp$ constant	$\asymp$ constant	$\asymp \ln n$
$\mathbb{Z}^k$	$\asymp n^k$	$\asymp n^{-k/2}$	$\asymp n$
$BS(1, N)$	$\asymp e^n$	$\asymp e^{-n^{1/3}}$	$\asymp n^{3/2}$
$\mathbb{Z} \wr \mathbb{Z}$	$n^n$	$\asymp  S ^{-n^{1/3}(\ln n)^{2/3}}$	$\asymp \ln(n)n^{3/2}$
$K \wr \mathbb{Z}$	$f(n)^n$	$\asymp e^{-n^{1/2}}$	$\asymp n^2$

$K$  polycyclic with exponential growth, and Følner function  $f(n)$ .

The groups  $BS(1, N) = \langle a, t \mid tat^{-1}a^{-N} \rangle$  are amenable groups.

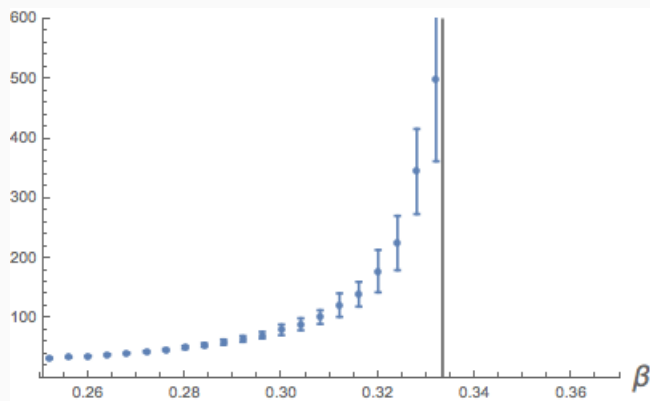
The groups  $BS(1, N) = \langle a, t \mid tat^{-1}a^{-N} \rangle$  are amenable groups.

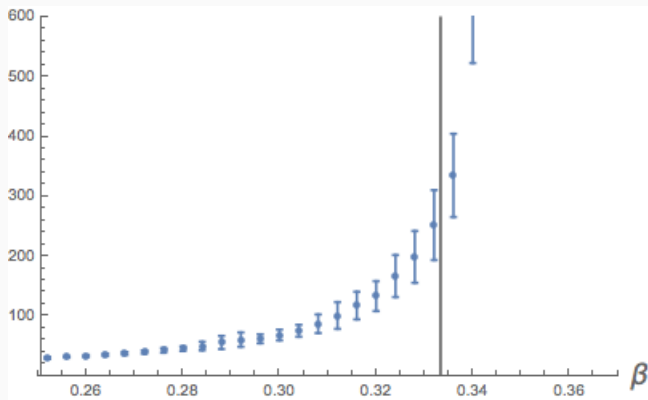
They approach  $\mathbb{Z} \wr \mathbb{Z}$  in the space of marked groups.

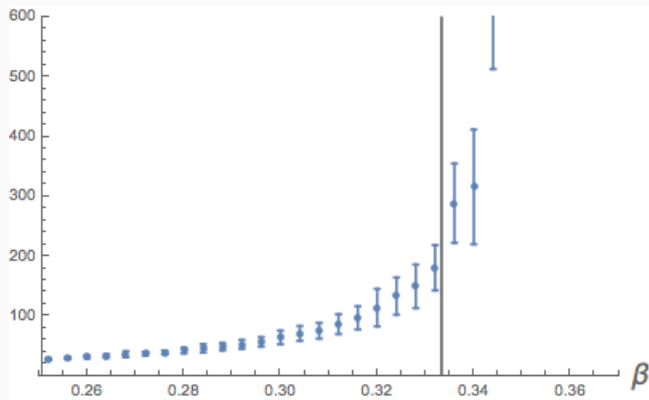
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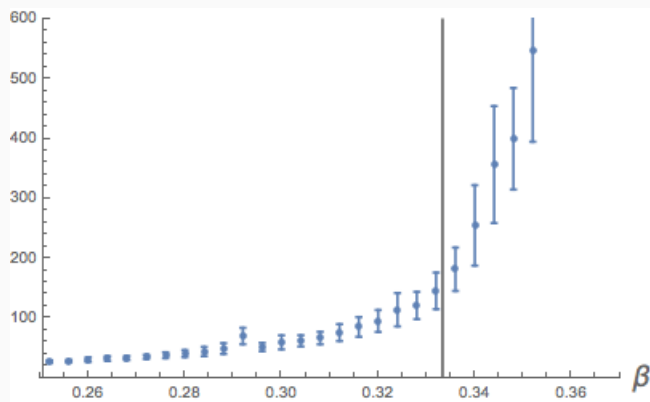
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So how does the ERR algorithm perform for  $BS(1, N)$  as  $N$  increases?

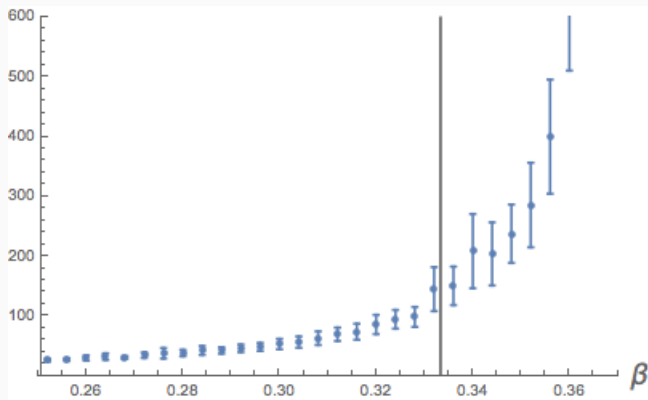












## WHAT IS GOING WRONG?

Just one relator, so not the problem discussed before.

We postulate that the cause of error in ERR comes from groups with bad sub-dominant behaviour of cogrowth function, which is measured by  $\mathcal{R}(n)$ .

Cameron to give details tomorrow. Details in [7]

THANK YOU



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