

The maximum of the periodogram of a sequence of functional data

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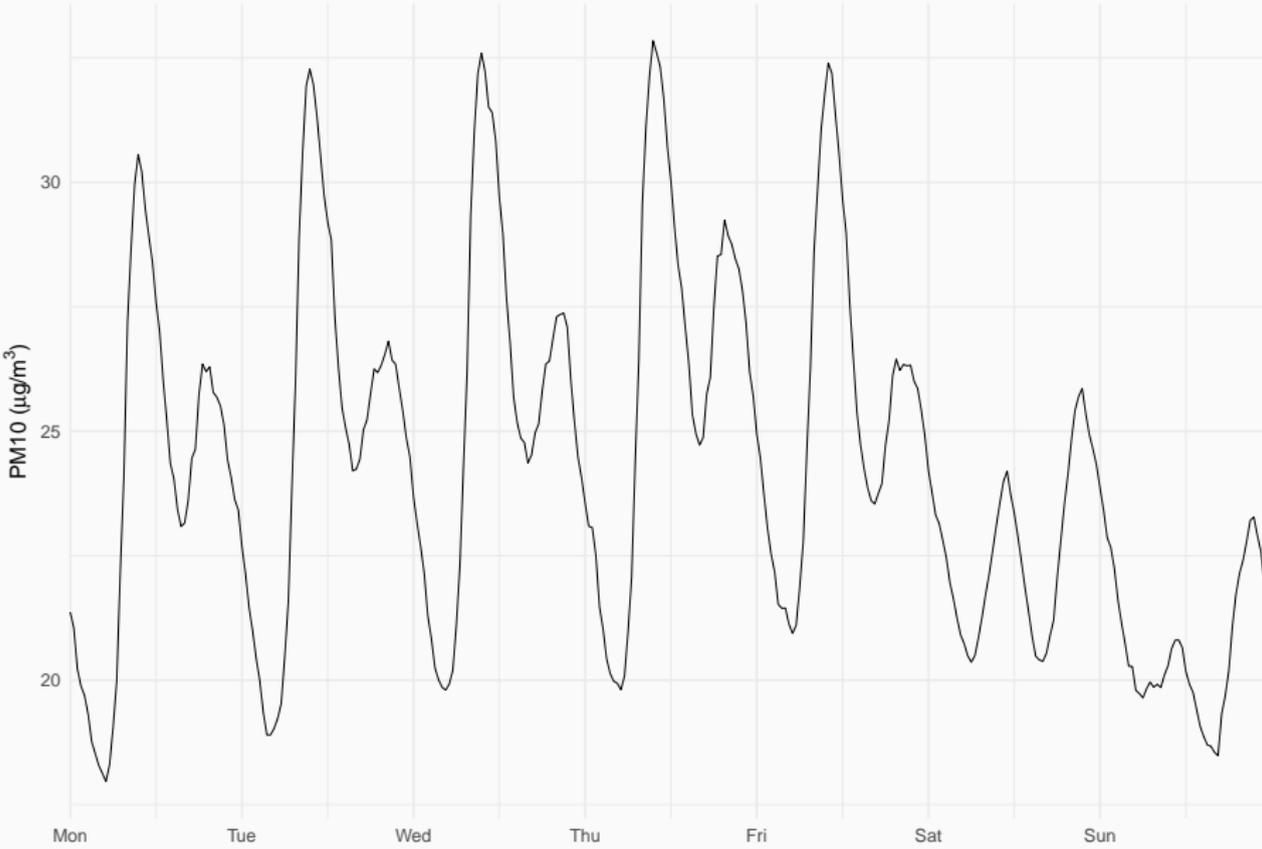
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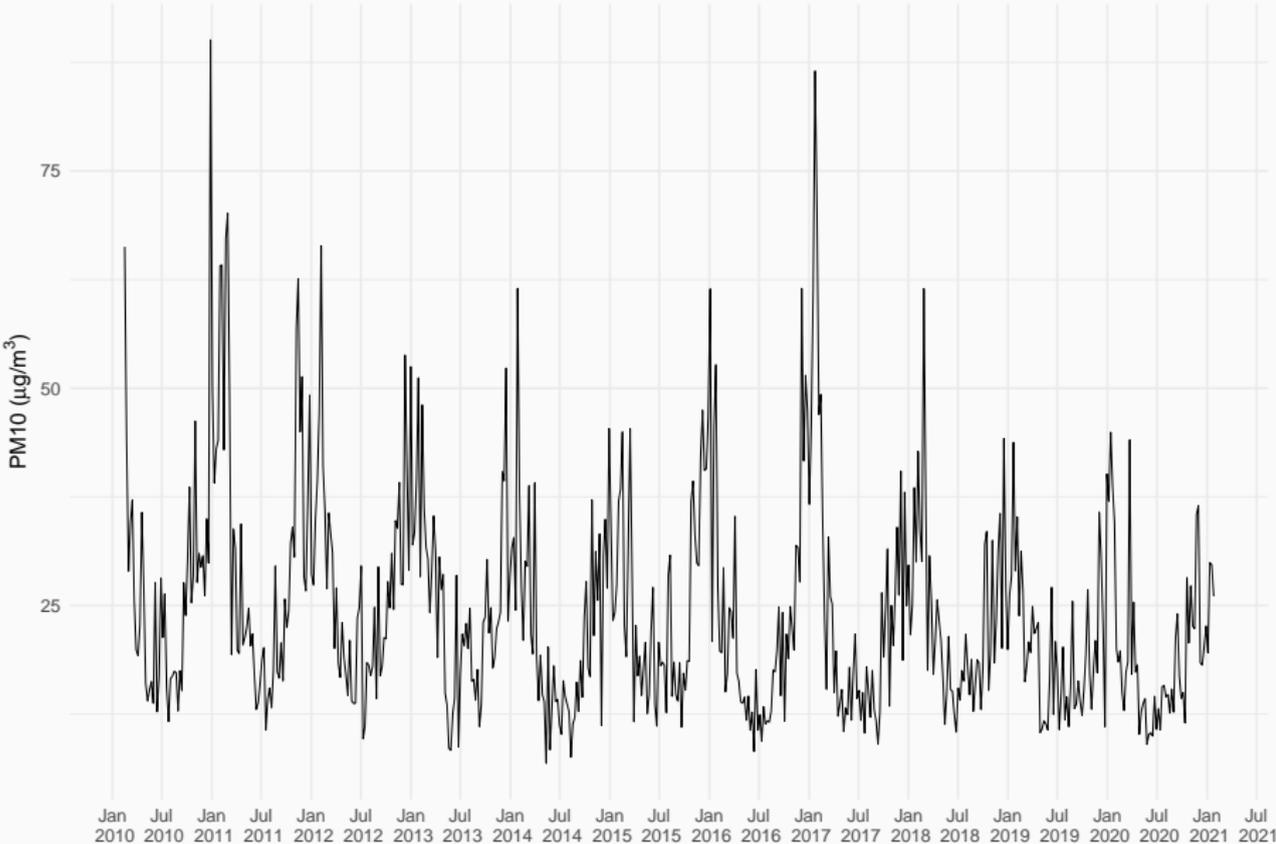
Motivation and problem

- Air quality data from Graz, Austria.
- The amount of particulate matter with a diameter of $10\ \mu\text{m}$ or less (PM10) is measured.
- PM10 can settle in the bronchi and lungs and cause health problems.
- Starting on February 18, 2010, the amount of PM10 in $\mu\text{g}/\text{m}^3$ is recorded every 30 minutes resulting in 48 observations per day.

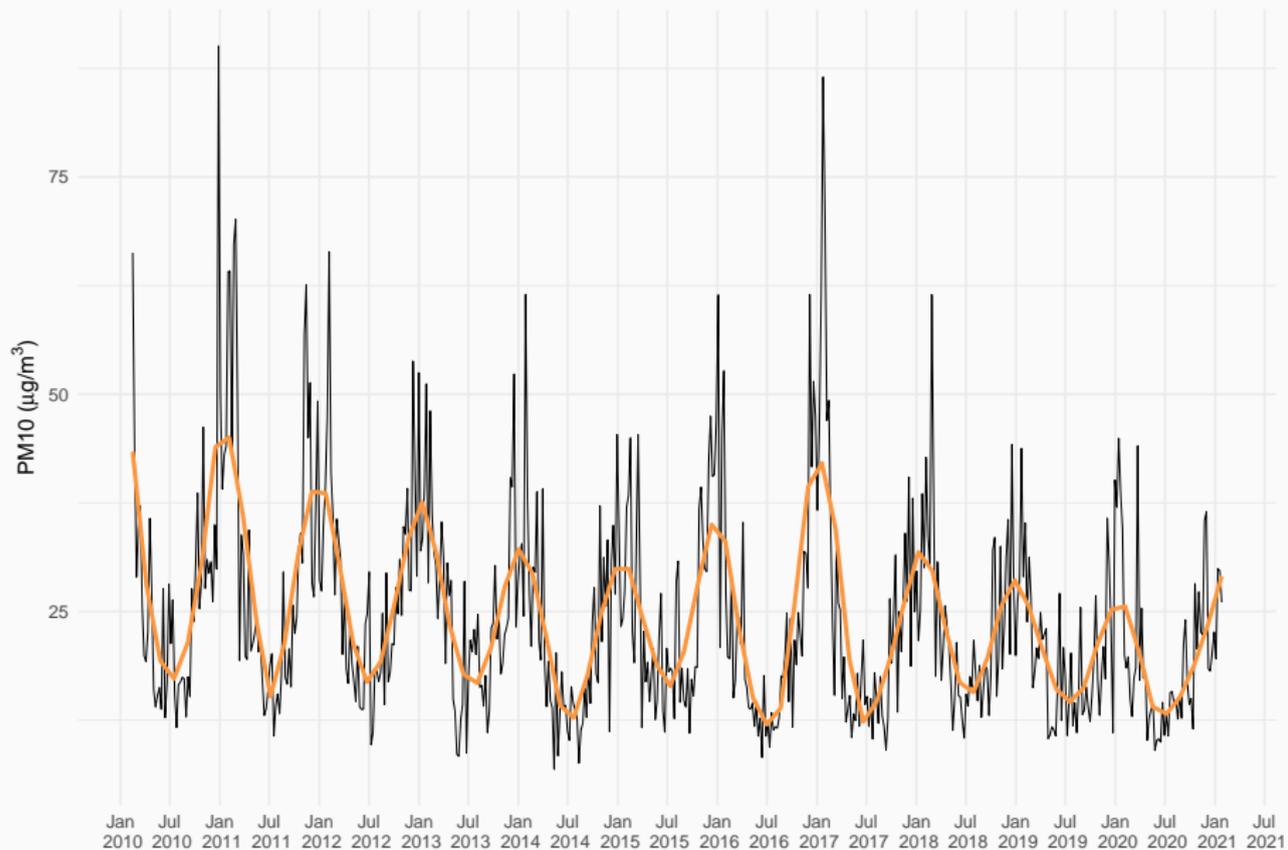
Weekly mean curve



Weekly averages



Weekly averages



Functional time series

- A functional time series is a sequence $\{X_t\}_{t \in \mathbb{Z}}$ such that each X_t is a curve $\{X_t(u)\}_{u \in [0,1]}$.
- We separate a continuous time process $\{\xi(u)\}_{u \in \mathbb{R}}$ using natural consecutive intervals, i.e.

$$X_t(u) = \xi(t + u)$$

for $u \in [0, 1]$ and $t \in \mathbb{Z}$.

- Such segmentation accounts for a periodic structure in the underlying continuous time process.
- There might still remain a periodic signal with respect to the discrete time parameter $t \in \mathbb{Z}$.

Model

$\{X_t\}_{t \in \mathbb{Z}}$ is a time series with values in a real separable Hilbert space \mathbb{H} (e.g. $L^2[0, 1]$) defined by

$$X_t = \mu + s_t + Y_t$$

for each $t \in \mathbb{Z}$, where

- $\mu \in \mathbb{H}$;
- $\{s_t\}_{t \in \mathbb{Z}} \subset \mathbb{H}$ is a deterministic sequence such that

$$s_t = s_{t+T} \quad \text{and} \quad \sum_{t=1}^T s_t = 0$$

for all $t \in \mathbb{Z}$ with some $T \geq 2$;

- $\{Y_t\}_{t \in \mathbb{Z}}$ is a stationary sequence of zero mean random elements with values in \mathbb{H} .

Hypothesis testing

We develop a methodology to test

$$H_0 : X_t = \mu + Y_t \quad \text{versus} \quad H_1 : X_t = \mu + S_t + Y_t$$

with an unknown $T \geq 2$.

Main results

Frequency domain approach

Our methodology is based on the frequency domain approach to the analysis of functional time series.

Definition

The discrete Fourier transform (DFT) of X_1, \dots, X_n is defined by

$$\mathcal{X}_n(\omega_j) = n^{-1/2} \sum_{t=1}^n X_t e^{-it\omega_j}$$

for $n \geq 1$, where

- i) $\omega_j = 2\pi j/n$ with $j = -\lfloor (n-1)/2 \rfloor, \dots, \lfloor n/2 \rfloor$ are the Fourier frequencies;
- ii) $i = \sqrt{-1}$.

Maximum of periodogram

The test statistic is given by

$$M_n = \max_{1 \leq j \leq q} \|\mathcal{X}_n(\omega_j)\|^2$$

for $n > 2$, where

- i) $\omega_j = 2\pi j/n$ with $1 \leq j \leq q = \lfloor n/2 \rfloor$;
- ii) $\|\cdot\|$ is the norm of the complexification of \mathbb{H} .

Maximum of periodogram

The test statistic is given by

$$M_n = \max_{1 \leq j \leq q} \|\mathcal{X}_n(\omega_j)\|^2$$

for $n > 2$.

- Small values of M_n indicate that there is no periodic component.
- Large values of M_n indicate that there is a periodic component.
- We need a criterion to decide when M_n is small and when M_n is large.

Linear processes

Suppose that $\{Y_t\}_{t \in \mathbb{Z}}$ is a linear process with values in \mathbb{H} given by

$$Y_t = \sum_{k=-\infty}^{\infty} a_k(\varepsilon_{t-k})$$

for each $t \in \mathbb{Z}$, where

- $\{\varepsilon_t\}_{t \in \mathbb{Z}}$ are iid zero mean random elements with values in \mathbb{H} ;
- $\{a_k\}_{k \in \mathbb{Z}} \subset L(\mathbb{H})$.

Assumptions

Assumption 1

- i) $E\|\varepsilon_0\|^r < \infty$ where $r > 2$ if $\dim \mathbb{H} < \infty$ and $r \geq 4$ otherwise;
- ii) the eigenvalues λ_k of $E[\varepsilon_0 \otimes \varepsilon_0]$ are distinct and the sequence $\{k\lambda_k\}_{k \geq 1}$ is ultimately non-increasing;
- iii) some technical conditions on the decay rate of $\{\lambda_k\}_{k \geq 1}$.

Assumption 2

- i) $\sum_{k \neq 0} \log(|k|) \|a_k\| < \infty$;
- ii) $A^{-1}(\omega)$ exists for each $\omega \in [-\pi, \pi]$, where $A(\omega) = \sum_{k=-\infty}^{\infty} a_k e^{-ik\omega}$ with $\omega \in [-\pi, \pi]$ is the transfer function;
- iii) $\sup_{\omega \in [0, \pi]} \|A^{-1}(\omega)\| < \infty$.

Theorem

Under H_0 and Assumptions 1 and 2, we have that

$$\lambda_1^{-1} \left(\max_{1 \leq j \leq q} \|A^{-1}(\omega_j) \mathcal{X}_n(\omega_j)\|^2 - b_n \right) \xrightarrow{d} G \quad \text{as } n \rightarrow \infty,$$

where

- $A(\omega_j) = \sum_{k=-\infty}^{\infty} a_k e^{-ik\omega_j}$ with $j = 1, \dots, q$;
- $b_n = \lambda_1 \log q - \lambda_1 \sum_{j=2}^{\infty} \log(1 - \lambda_j/\lambda_1)$;
- G is the standard Gumbel distribution with the CDF given by $F(x) = \exp\{-\exp\{-x\}\}$ for $x \in \mathbb{R}$.

$\{Y_t\}_{t \in \mathbb{Z}}$ is an FAR(1) model given by

$$Y_t = \rho(Y_{t-1}) + \varepsilon_t = \sum_{j=0}^{\infty} \rho^j(\varepsilon_{t-j})$$

for $t \in \mathbb{Z}$ with $\rho \in L(\mathbb{H})$.

Assumption 3

- i) There is an $n_0 \geq 1$ such that $\|\rho^{n_0}\| < 1$;
- ii) $\hat{\rho}$ is an estimator of ρ such that

$$\|\hat{\rho} - \rho\|_{op} = o_p(1/\tau'_n)$$

as $n \rightarrow \infty$ with $\tau'_n \geq \log n$.

The transfer function, residuals and their eigenvalues

- $\{\hat{\varepsilon}_k\}_{2 \leq k \leq n}$ are the residuals given by

$$\hat{\varepsilon}_k = X_k - \hat{\rho}(X_{k-1})$$

for $k = 2, \dots, n$.

- $\{\hat{\lambda}_j\}_{j \geq 1}$ are the eigenvalues of

$$\frac{1}{n-1} \sum_{k=2}^n \hat{\varepsilon}_k \otimes \hat{\varepsilon}_k.$$

- The transfer function $A(\omega) = (I - e^{-i\omega} \rho)^{-1}$ and hence $A^{-1}(\omega) = I - e^{-i\omega} \rho$ for $\omega \in [-\pi, \pi]$.

Theorem

Under H_0 and Assumptions 1 and 3,

$$G_n := \hat{\lambda}_1^{-1} \max_{1 \leq j \leq q} \|(I - e^{-i\omega_j} \hat{\rho})(\mathcal{X}_n(\omega_j))\|^2$$
$$- \log q + \max \left\{ \sum_{j=2}^{\tau_n} \log(1 - \hat{\lambda}_j / \hat{\lambda}_1), c_n \right\} \xrightarrow{d} \mathcal{G}$$

as $n \rightarrow \infty$, where $\{\tau_n\}_{n \geq 1} \subset \mathbb{N}$ and $\{c_n\}_{n \geq 1} \subset \mathbb{R}$ are sequences that satisfy certain technical conditions.

Theorem

Under H_1 ,

$$G_n/\ell_n \xrightarrow{P} \infty \quad \text{as } n \rightarrow \infty$$

for any positive sequence $\ell_n = o(n)$ as $n \rightarrow \infty$ provided certain technical conditions are satisfied.

Empirical study

- We plot the points $(j, G_n(j))$ with $j = 1, \dots, q = 1998$ and

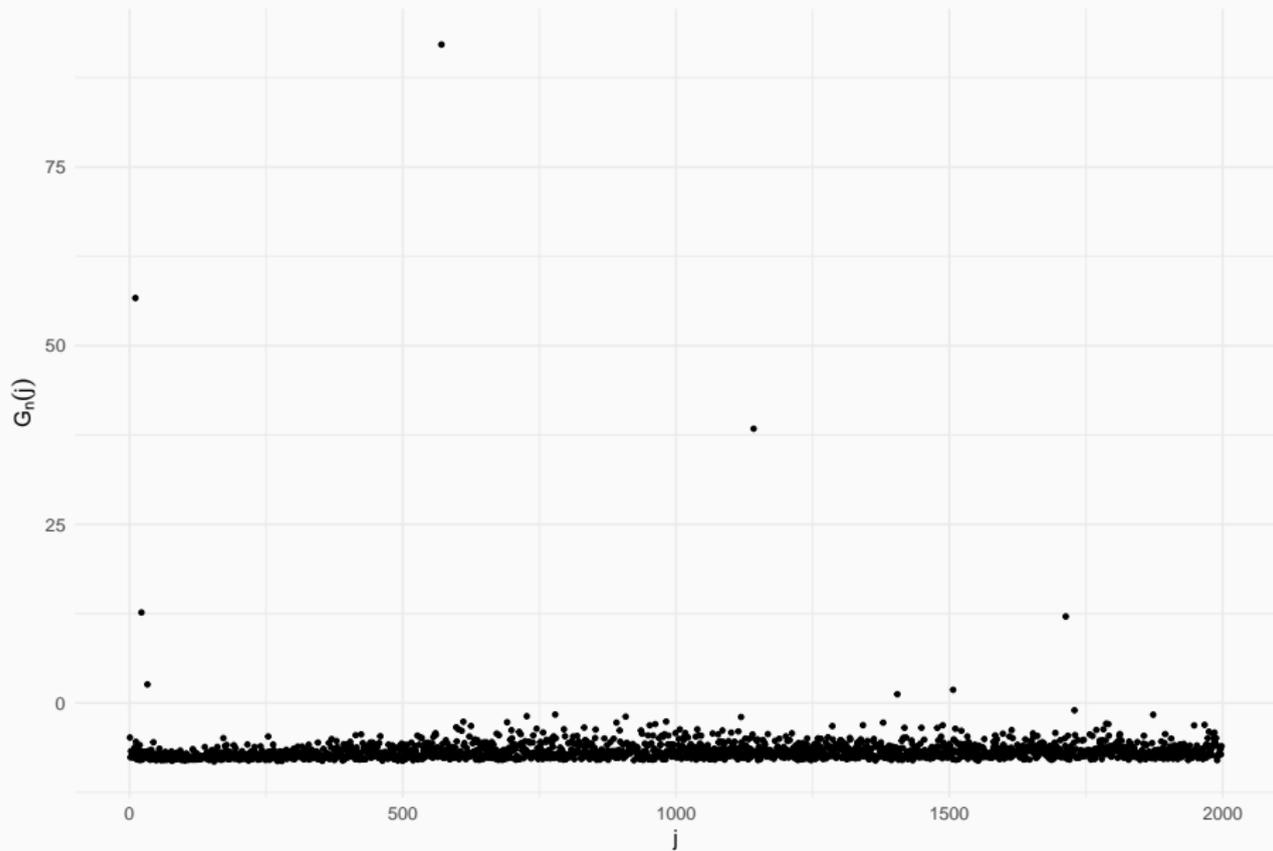
$$G_n(j) := \lambda_1^{-1} \|(I - e^{-i\omega_j} \hat{\rho})(\mathcal{X}_n(\omega_j))\|^2 - \log q + \max \left\{ \sum_{j=2}^{\tau_n} \log(1 - \hat{\lambda}_j / \hat{\lambda}_1), c_n \right\},$$

where $n = 3997$.

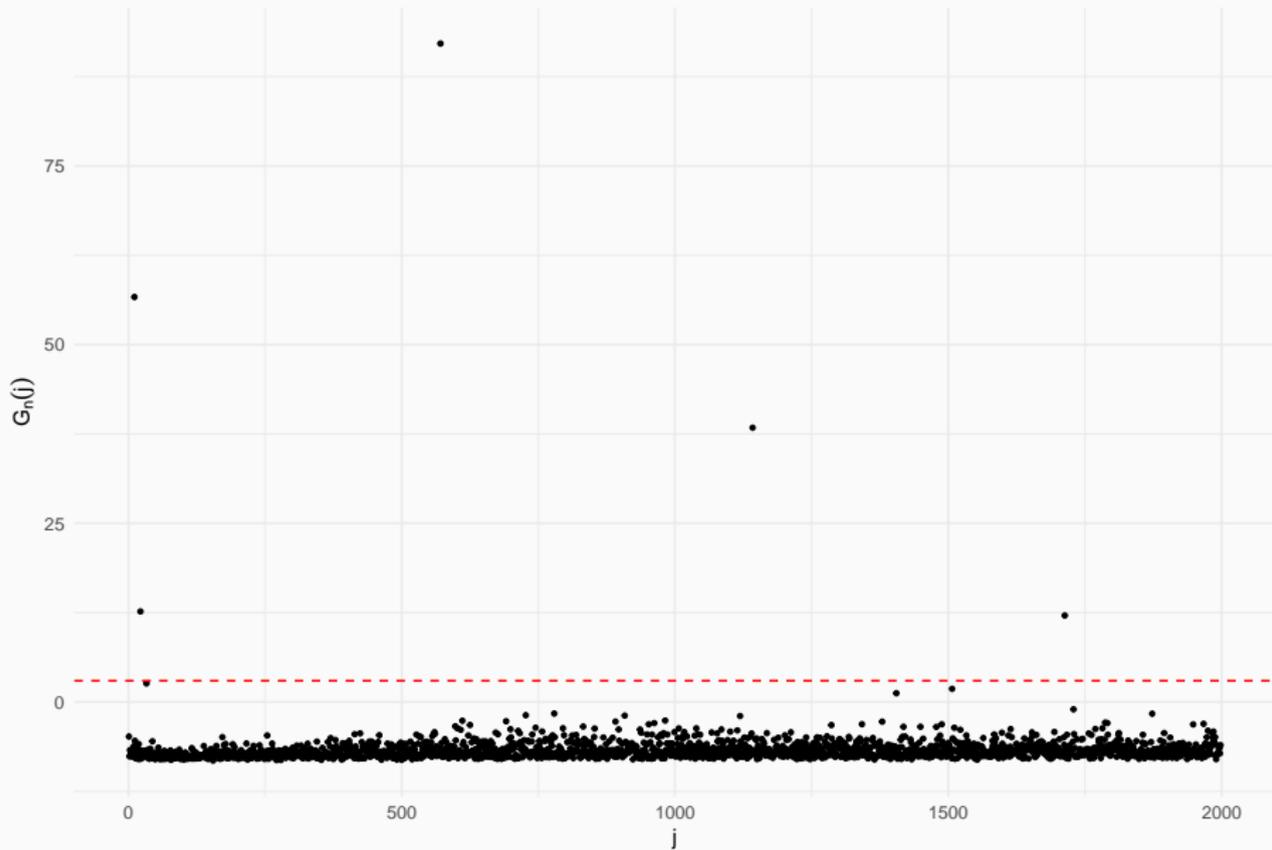
- Observe that

$$G_n = \max_{1 \leq j \leq q} G_n(j).$$

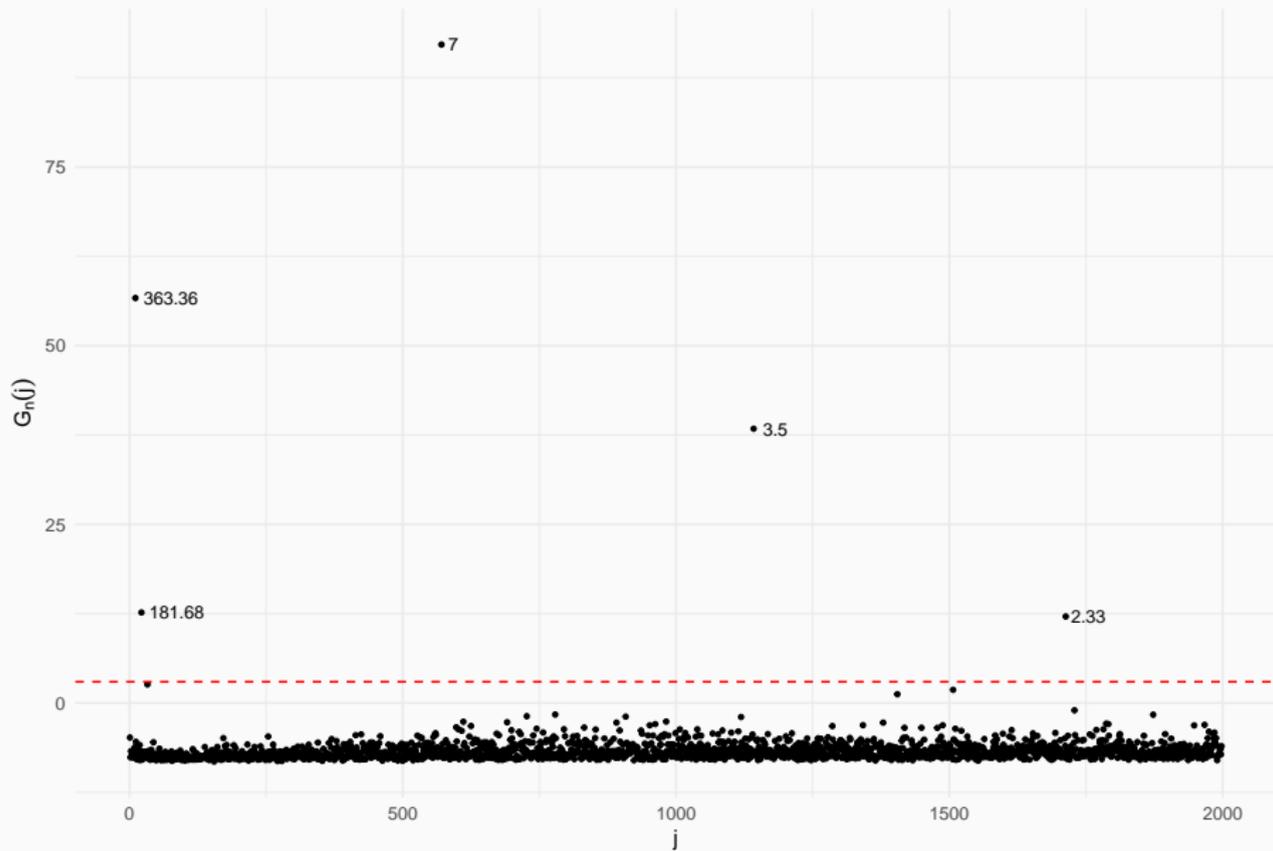
PM10 time series



PM10 time series



PM10 time series



Representation of periodic signals

Lemma

Suppose that $\{s_t\}_{t \in \mathbb{Z}}$ is a deterministic sequence with values in \mathbb{H} such that

$$s_t = s_{t+T} \quad \text{and} \quad \sum_{t=1}^T s_t = 0$$

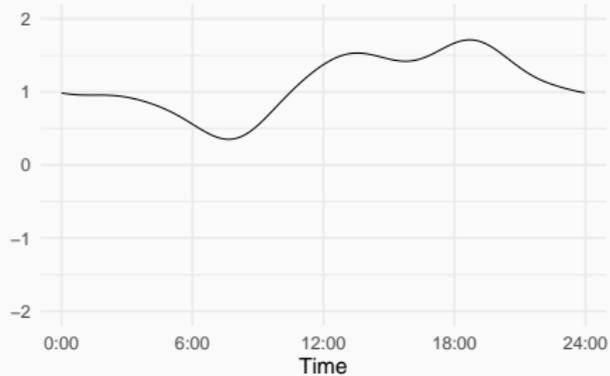
for all $t \in \mathbb{Z}$ with some $T \geq 2$. Then there exist $w_{11}, \dots, w_{1\lfloor T/2 \rfloor} \in \mathbb{H}$ and $w_{21}, \dots, w_{2\lfloor T/2 \rfloor} \in \mathbb{H}$ such that

$$s_t = \sum_{k=1}^{\lfloor T/2 \rfloor} [\cos(2\pi kt/T)w_{1k} + \sin(2\pi kt/T)w_{2k}]$$

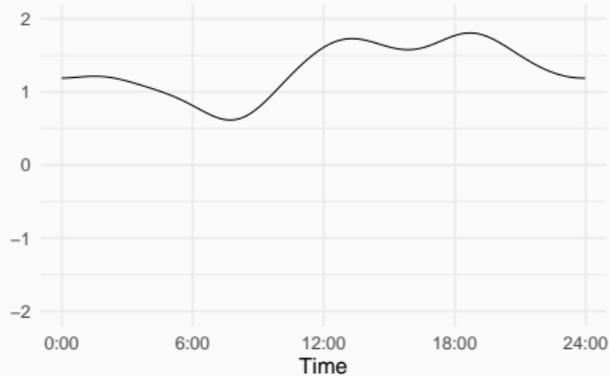
for all $t \in \mathbb{Z}$.

Yearly periodic component

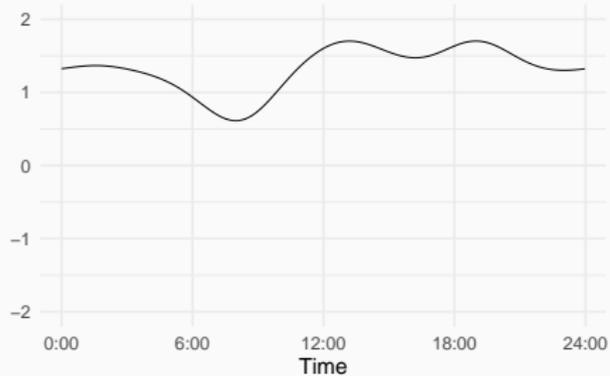
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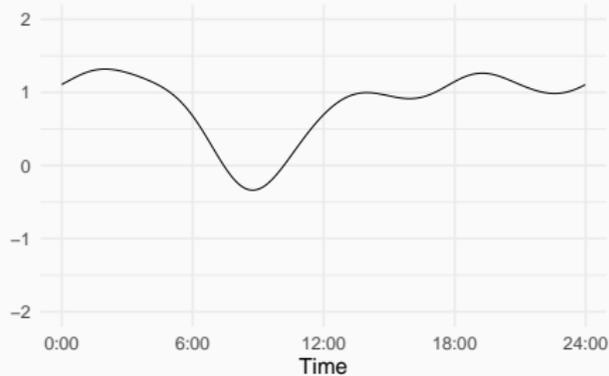
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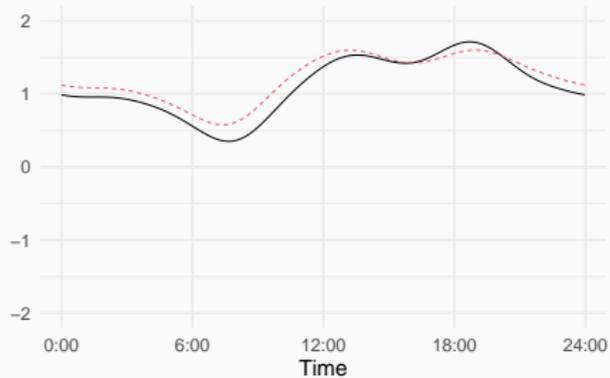


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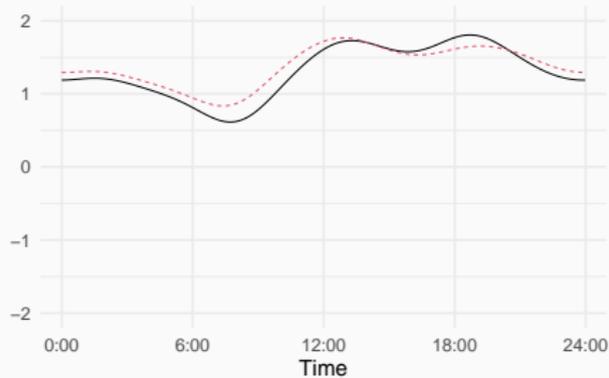


Periodic component

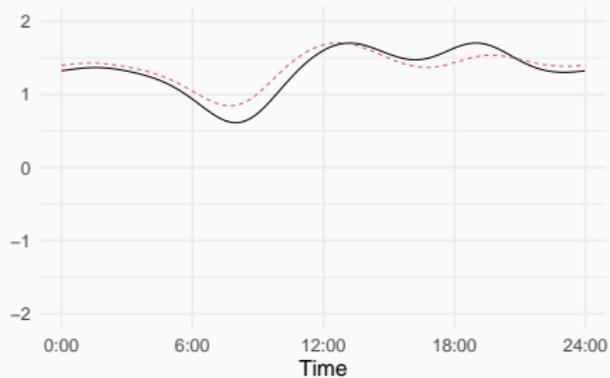
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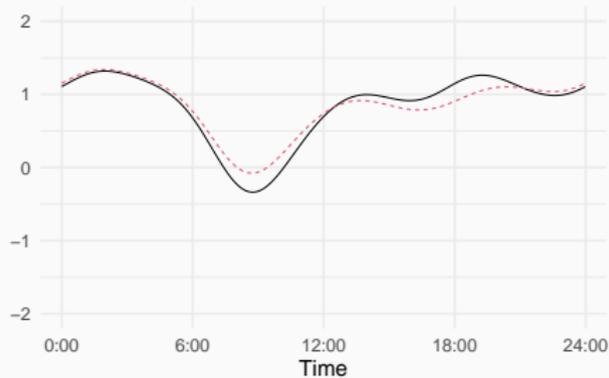
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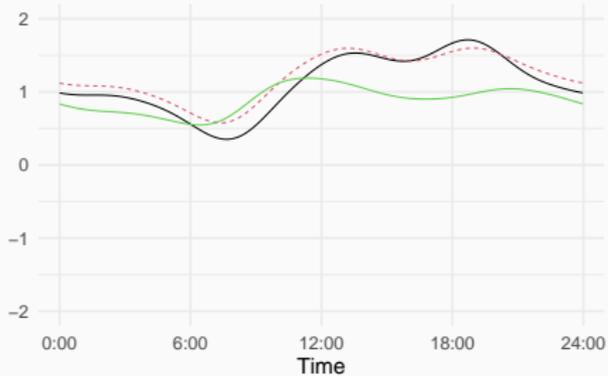


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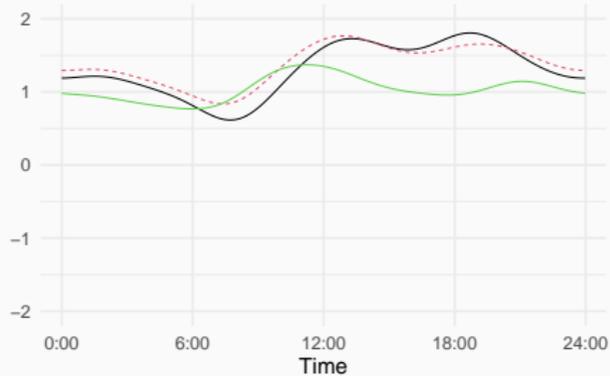


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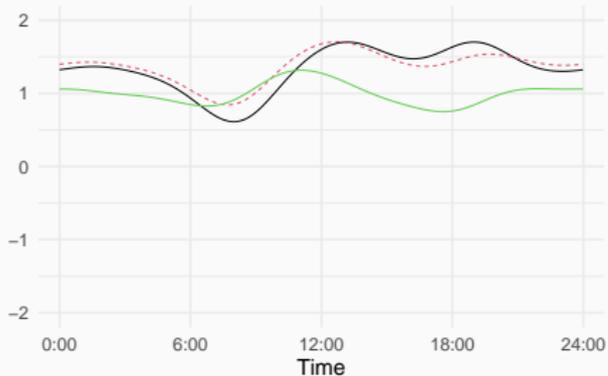
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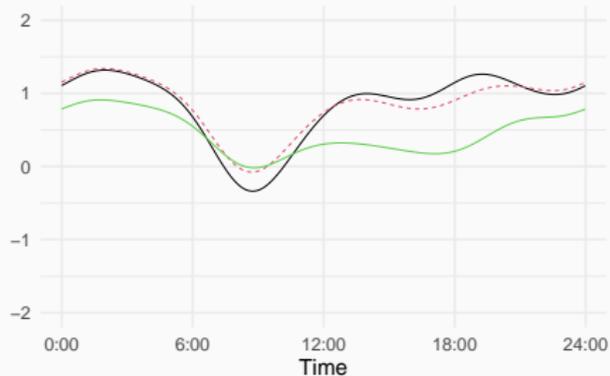
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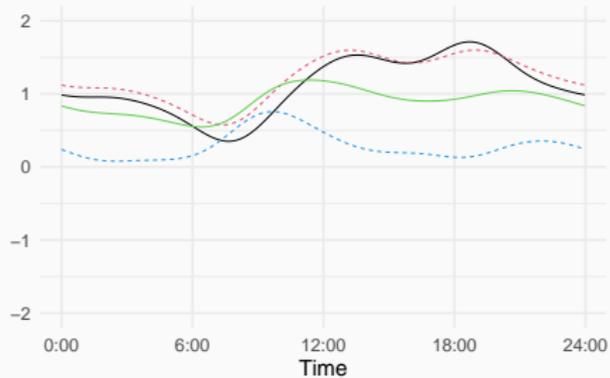


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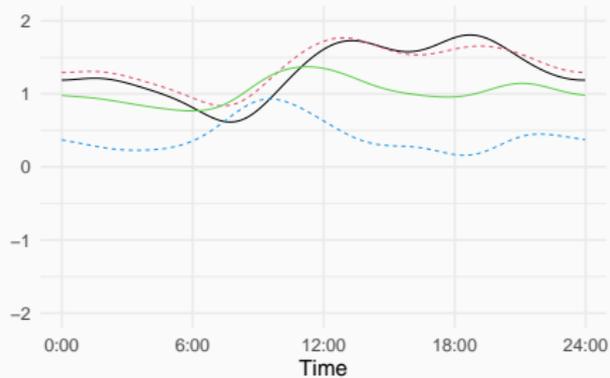


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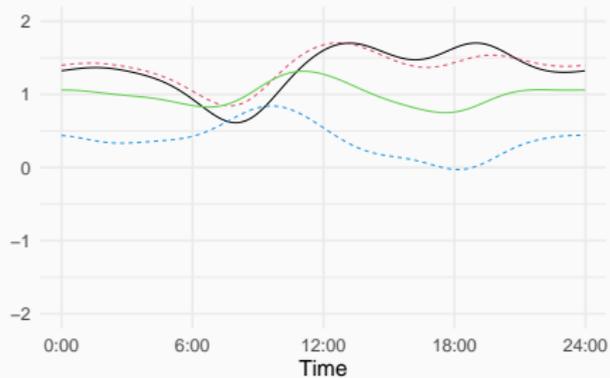
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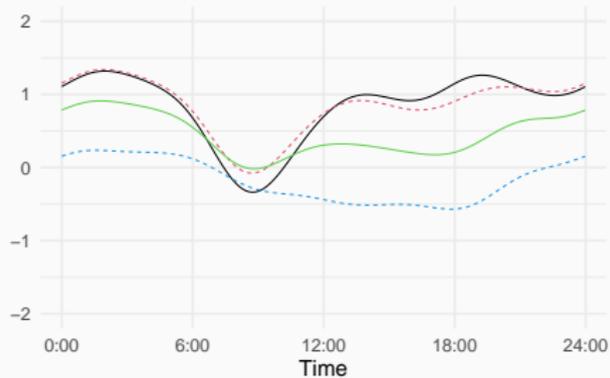
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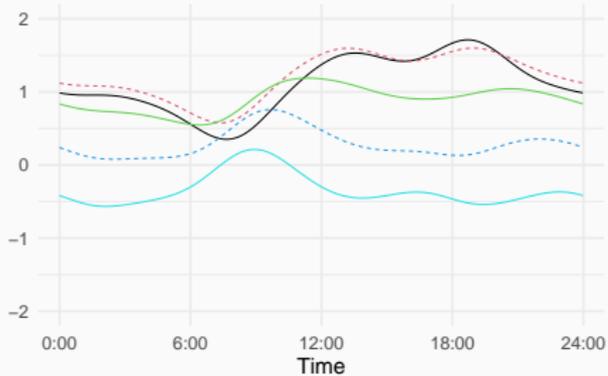


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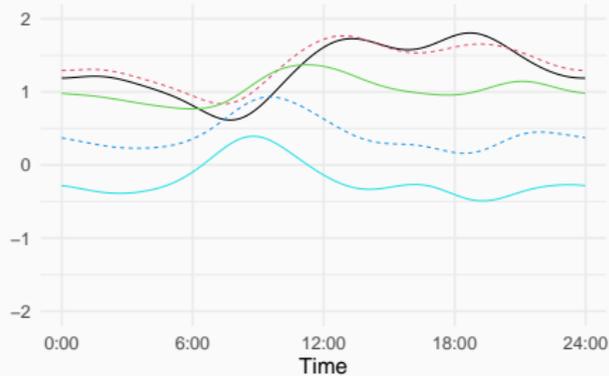


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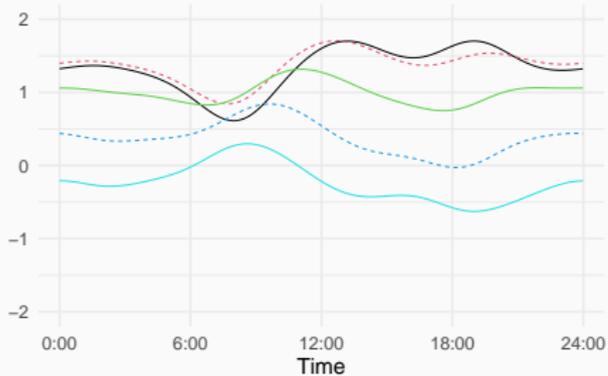
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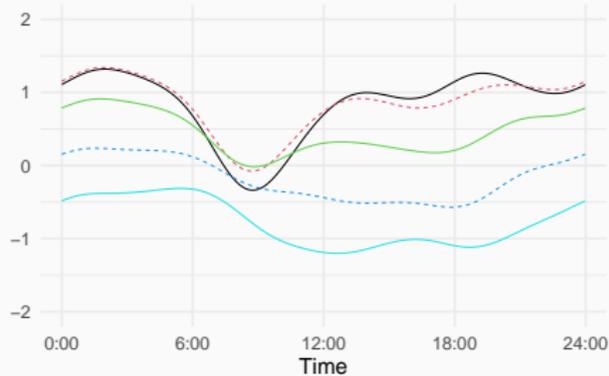
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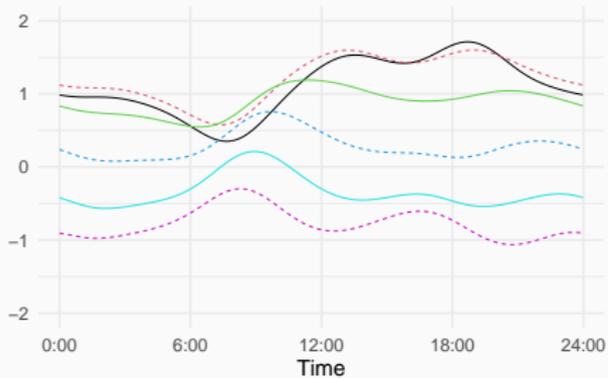


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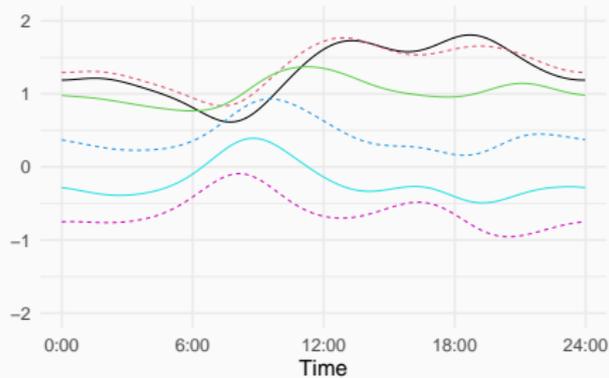


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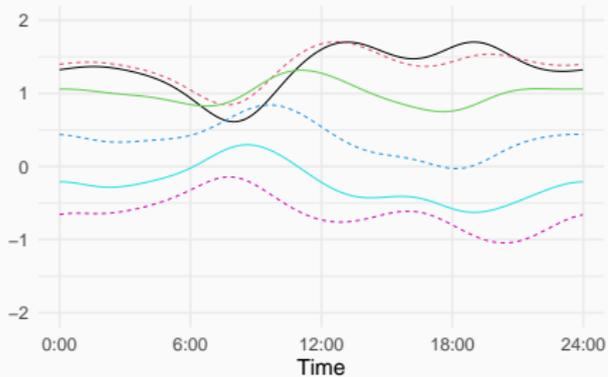
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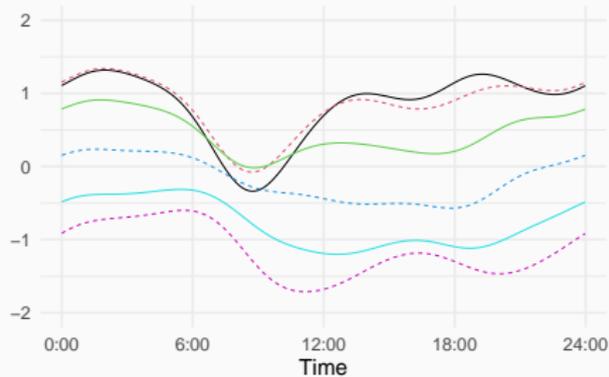
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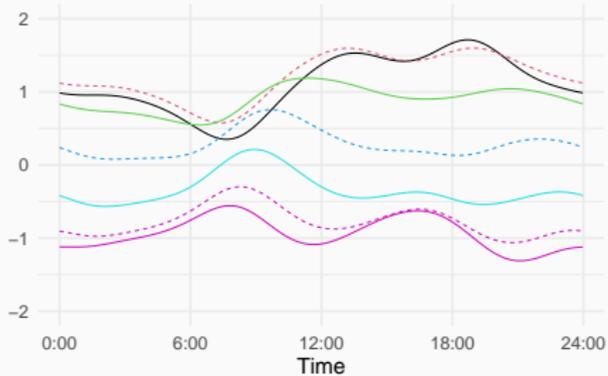


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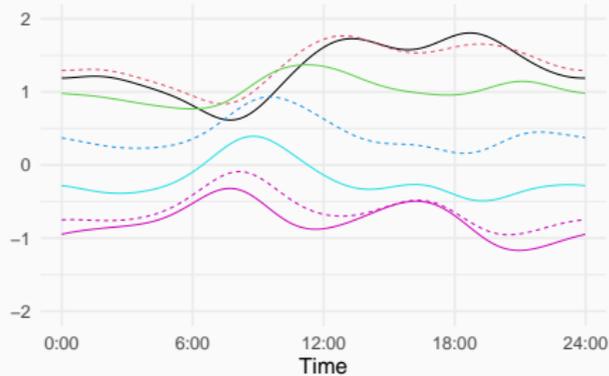


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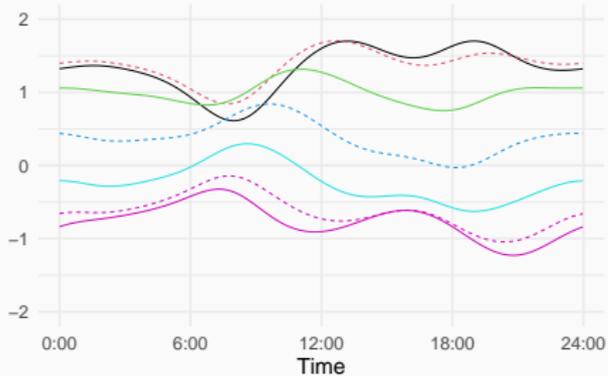
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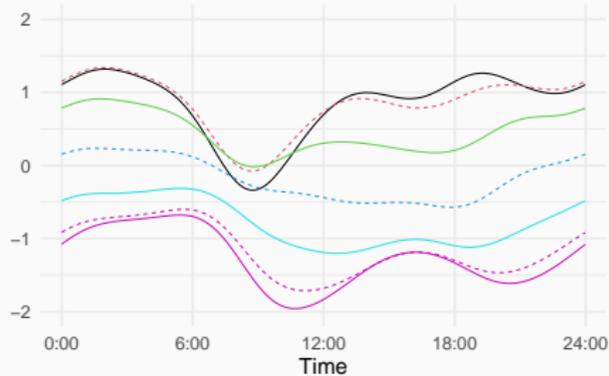
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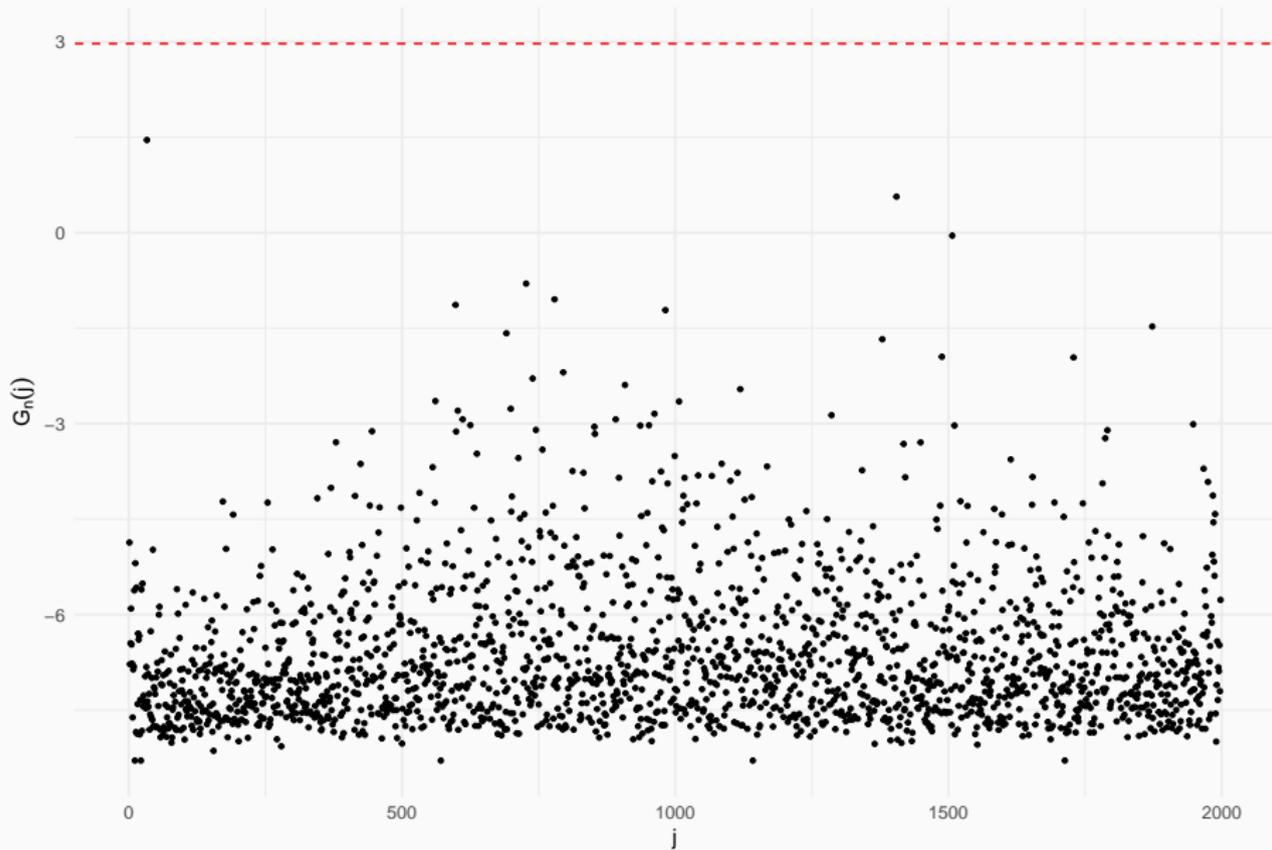
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Sunday



Deseasonalized data



Summary

Summary

- A general test for periodic signals in Hilbert space valued time series when the length of the period is unknown.
- The appropriately standardized maximum of the periodogram converges in distribution to the standard Gumbel distribution.
- A weekly as well as a yearly periodic components are detected in the PM10 data.
- The periodic signals in the PM10 data are not pure sinusoids but are actually driven by several sinusoids.

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