## The maximum of the periodogram of a sequence of functional data

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Motivation and problem

## PM10 data

- Air quality data from Graz, Austria.
- The amount of particulate matter with a diameter of $10 \mu \mathrm{~m}$ or less (PM10) is measured.
- PM10 can settle in the bronchi and lungs and cause health problems.
- Starting on February 18, 2010, the amount of PM10 in $\mu \mathrm{g} / \mathrm{m}^{3}$ is recorded every 30 minutes resulting in 48 observations per day.


## Raw data



## Weekly mean curve



## Weekly averages



## Weekly averages



## The PM10 data as a sequence of curves

We investigate the PM10 data as a functional time series, i.e. as a sequence of daily curves.

## Model

$\left\{X_{t}\right\}_{t \in \mathbb{Z}}$ is a time series with values in a real separable Hilbert space $\mathbb{H}$ defined by

$$
X_{t}=\mu+s_{t}+Y_{t}
$$

for each $t \in \mathbb{Z}$, where

- $\mu \in \mathbb{H} ;$
- $\left\{s_{t}\right\}_{t \in \mathbb{Z}} \subset \mathbb{H}$ is a deterministic sequence such that

$$
S_{t}=S_{t+T} \quad \text { and } \quad \sum_{t=1}^{T} s_{t}=0
$$

for all $t \in \mathbb{Z}$ with some $T \geq 2$;

- $\left\{Y_{t}\right\}_{t \in \mathbb{Z}}$ is a stationary sequence of zero mean random elements with values in $\mathbb{H}$.


## Hypothesis testing

We develop a methodology to test

$$
H_{0}: X_{t}=\mu+Y_{t} \quad \text { versus } \quad H_{1}: X_{t}=\mu+S_{t}+Y_{t}
$$

with an unknown $T \geq 2$.

Main results

## Frequency domain approach

Our methodology is based on the frequency domain approach to the analysis of functional time series.

## DFT

## Definition

The discrete Fourier transform (DFT) of $X_{1}, \ldots, X_{n}$ is defined by

$$
\mathcal{X}_{n}\left(\omega_{j}\right)=n^{-1 / 2} \sum_{t=1}^{n} x_{t} e^{-i t \omega_{j}}
$$

for $n \geq 1$, where $\omega_{j}=2 \pi j / n$ with $j=-\lfloor(n-1) / 2\rfloor, \ldots,\lfloor n / 2\rfloor$ and $i=\sqrt{-1}$.

## Maximum of periodogram

The test statistic is given by

$$
\max _{1 \leq j \leq q}\left\|\mathcal{X}_{n}\left(\omega_{j}\right)\right\|^{2}
$$

for $n>1$, where
i) $\omega_{j}=2 \pi j / n$ with $1 \leq j \leq q=\lfloor n / 2\rfloor$;
ii) $\|\cdot\|$ is the norm of the complexification of $\mathbb{H}$.

## Linear processes

Suppose that $\left\{Y_{t}\right\}_{t \in \mathbb{Z}}$ is a linear process with values in $\mathbb{H}$ given by

$$
Y_{t}=\sum_{k=-\infty}^{\infty} a_{k}\left(\varepsilon_{t-k}\right)
$$

for each $t \in \mathbb{Z}$, where

- $\left\{a_{k}\right\}_{k \in \mathbb{Z}} \subset L(\mathbb{H}) ;$
- $\left\{\varepsilon_{t}\right\}_{t \in \mathbb{Z}}$ are iid zero mean random elements with values in $\mathbb{H}$.


## Assumptions

## Assumption 1

i) $E\left\|\varepsilon_{0}\right\|^{r}<\infty$ where $r>2$ if $\operatorname{dim} \mathbb{H}<\infty$ and $r \geq 4$ otherwise;
ii) the eigenvalues $\lambda_{k}$ of $E\left[\varepsilon_{0} \otimes \varepsilon_{0}\right]$ are distinct and the sequence $\left\{k \lambda_{k}\right\}_{k \geq 1}$ is ultimately non-increasing;
iii) some technical conditions on the decay rate of $\left\{\lambda_{k}\right\}_{k \geq 1}$.

## Assumption 2

i) $\sum_{k \neq 0} \log (|k|)\left\|a_{k}\right\|_{o p}<\infty$;
ii) $A^{-1}(\omega)$ exists for each $\omega \in[-\pi, \pi]$, where $A(\omega)=\sum_{k=-\infty}^{\infty} a_{k} e^{-i k \omega}$ with $\omega \in[-\pi, \pi]$;
iii) $\sup _{\omega \in[0, \pi]}\left\|A^{-1}(\omega)\right\|_{o p}<\infty$.

## Main result

## Theorem

Under $\mathrm{H}_{0}$ and Assumptions 1 and 2, we have that

$$
\lambda_{1}^{-1}\left(\max _{1 \leq j \leq q}\left\|A^{-1}\left(\omega_{j}\right) \mathcal{X}_{n}\left(\omega_{j}\right)\right\|^{2}-b_{n}\right) \xrightarrow{d} G \quad \text { as } \quad n \rightarrow \infty,
$$

where

- $A\left(\omega_{j}\right)=\sum_{k=-\infty}^{\infty} a_{k} e^{-i k \omega_{j}}$ with $j=1, \ldots, q$;
- $b_{n}=\lambda_{1} \log q-\lambda_{1} \sum_{j=2}^{\infty} \log \left(1-\lambda_{j} / \lambda_{1}\right)$;
- $G$ is the standard Gumbel distribution with the CDF given by $F(x)=\exp \{-\exp \{-x\}\}$ for $x \in \mathbb{R}$.


## High-dimensional Gaussian approximation

The core part of the proof is a high-dimensional Gaussian approximation for the DFT developed by Chernozhukov et al. (2017).

## FAR(1)

$\left\{Y_{t}\right\}_{t \in \mathbb{Z}}$ is an $\operatorname{FAR}(1)$ model given by

$$
Y_{t}=\rho\left(Y_{t-1}\right)+\varepsilon_{t}=\sum_{j=0}^{\infty} \rho^{j}\left(\varepsilon_{t-j}\right)
$$

for $t \in \mathbb{Z}$ with $\rho \in L(\mathbb{H})$.

## Assumption 3

i) There is an $n_{0} \geq 1$ such that $\left\|\rho^{n_{0}}\right\|<1$;
ii) $\hat{\rho}$ is an estimator of $\rho$ such that

$$
\|\hat{\rho}-\rho\|_{o p}=o_{p}\left(1 / \tau_{n}^{\prime}\right)
$$

as $n \rightarrow \infty$ with $\tau_{n}^{\prime} \geq \log n$.

## Residuals and their eigenvalues

- $\left\{\hat{\varepsilon}_{k}\right\}_{2 \leq k \leq n}$ are the residuals given by

$$
\hat{\varepsilon}_{k}=X_{k}-\hat{\rho}\left(X_{k-1}\right)
$$

for $k=2, \ldots, n$.

- $\left\{\hat{\lambda}_{j}\right\}_{j \geq 1}$ are the eigenvalues of

$$
\frac{1}{n-1} \sum_{k=2}^{n} \hat{\varepsilon}_{k} \otimes \hat{\varepsilon}_{k} .
$$

## Test statistic

## Theorem

Under $\mathrm{H}_{0}$ and Assumptions 1 and 3,

$$
G_{n}:=\hat{\lambda}_{1}^{-1} \max _{1 \leq j \leq q}\left\|\left(I-e^{-i \omega_{j}} \hat{\rho}\right)\left(\mathcal{X}_{n}\left(\omega_{j}\right)\right)\right\|^{2}
$$

$$
-\log q+\max \left\{\sum_{j=2}^{\tau_{n}} \log \left(1-\hat{\lambda}_{j} / \hat{\lambda}_{1}\right), c_{n}\right\} \xrightarrow{d} \mathcal{G}
$$

as $n \rightarrow \infty$, where $\left\{\tau_{n}\right\}_{n \geq 1} \subset \mathbb{N}$ and $\left\{c_{n}\right\}_{n \geq 1} \subset \mathbb{R}$ are sequences that satisfy certain technical conditions.

## Consistency

Theorem
Under $H_{1}$,

$$
G_{n} / \ell_{n} \xrightarrow{p} \infty \quad \text { as } \quad n \rightarrow \infty
$$

for any positive sequence $\ell_{n}=o(n)$ as $n \rightarrow \infty$ provided certain technical conditions are satisfied.

Empirical study

## PM10 time series

- We plot the points $\left(j, G_{n}(j)\right)$ with $j=1, \ldots, q=1998$ and

$$
\begin{aligned}
& G_{n}(j):=\lambda_{1}^{-1}\left\|\left(I-e^{-i \omega_{j}} \hat{\rho}\right)\left(\mathcal{X}_{n}\left(\omega_{j}\right)\right)\right\|^{2} \\
&-\log q+\max \left\{\sum_{j=2}^{\tau_{n}} \log \left(1-\hat{\lambda}_{j} / \hat{\lambda}_{1}\right), c_{n}\right\},
\end{aligned}
$$

where $n=3997$.

- Observe that

$$
G_{n}=\max _{1 \leq j \leq q} G_{n}(j) .
$$

## PM10 time series

## PM10 time series

75

50
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## PM10 time series

## Yearly periodic component



Friday


## Periodic component



Friday


## Periodic component

Monday


Friday


Wednesday


Sunday


## Periodic component

Monday


Friday


Wednesday



## Periodic component

Monday


Friday


Wednesday



## Periodic component

Monday


Friday


Wednesday



## Periodic component

Monday


Friday


Wednesday


Sunday


## Deseasonalized data



## Summary

## Summary

- A general test for periodic signals in Hilbert space valued time series when the length of the period is unknown.
- The appropriately standardized maximum of the periodogram converges in distribution to the standard Gumbel distribution.
- A weekly as well as a yearly periodic components are detected in the PM10 data.
- The periodic signals in the PM10 data are not pure sinusoids but are actually driven by several sinusoids.
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