

Data  $x_1 \dots x_n$

$$\text{Mean } \langle x \rangle = \frac{1}{n} \sum_{i=1}^n x_i$$

$$\text{Variance } \sigma^2 = \frac{1}{n} \sum_{i=1}^n (x_i - \langle x \rangle)^2$$

Std. deviation  $\sigma$

mean centering / standardisation

$$x_i' = \frac{x_i - \langle x \rangle}{\sigma}$$

$\leadsto$  mean 0, std dev of one

variation of zero-mean data  $\sigma^2 = \frac{1}{n} \sum_i x_i^2$

Covariance of two zero-mean measurements  
 $A = a_1, \dots, a_n$  and  $B = b_1, \dots, b_n$

$$\sigma_{AB} = \frac{1}{n} \sum_i a_i b_i$$

if positive  $\leadsto$  correlated data

if negative  $\leadsto$  negat. correlated data

if  $\emptyset$   $\leadsto$  uncorrelated

vector notation

$$a = (a_1 \dots a_n)$$

$$b = (b_1 \dots b_n)$$

$$\sigma_{ab}^2 = \frac{1}{n} ab^T$$

Data given as  $X = \begin{pmatrix} x_1 \\ \vdots \\ x_n \end{pmatrix}$  ( $m \times n$  matrix)

Covariance matrix  $C_x = \frac{1}{n} X X^T$  ( $m \times m$  matrix)

- symmetric
- diagonal elements indicate variances
- off-diagonal elements are co-variances

↪ example in  $\mathbb{R}$

# Principal Component Analysis

Intro ~ Slides

Note: now  $\vec{x}_i$  are columns, still a  $m \times n$  matrix

A basis  $B = \begin{pmatrix} \vec{b}_1 \\ \vec{b}_2 \\ \vdots \\ \vec{b}_m \end{pmatrix}$  our data  $X = (\vec{x}_1 \vec{x}_2 \dots \vec{x}_n)$   
 (I won't use  $\Rightarrow$ )

Change of basis via transformation  $P$

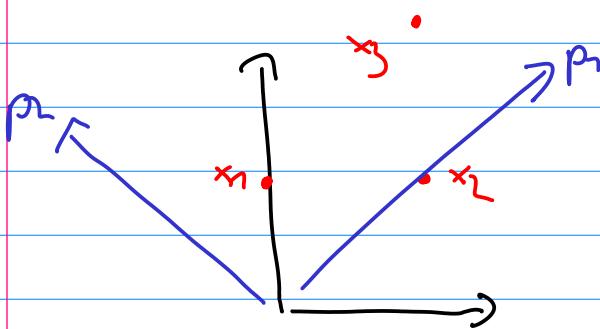
$P \cdot X = Y$   $P = \begin{pmatrix} p_1 \\ p_2 \\ \vdots \\ p_m \end{pmatrix}$   $X = (x_1 | x_2 | \dots | x_n)$   
 $m \times m \mid m \times n \mid m \times n$

$P$  transforms  $X$  into  $Y$  (rotation, stretch)

$p_1, \dots, p_m$  are a new basis

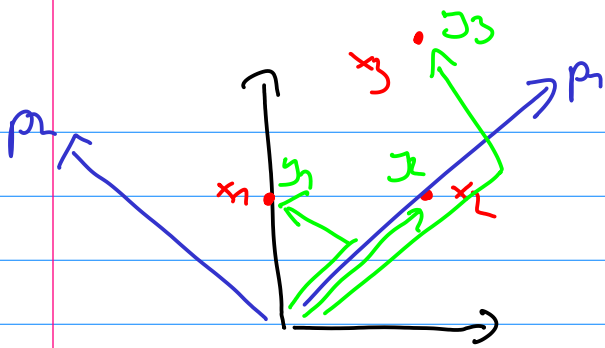
having unit lengths  
 i.e.  $\|p_i\| = 1$

stretchy example: ( $p_i$  are not (pairwise) orthogonal, only orthogonal, i.e.  $p_i \cdot p_j = 0$  for  $i \neq j$ )



$x_1 = \begin{pmatrix} 0 \\ 3 \end{pmatrix}$   $p_1 = \begin{pmatrix} \frac{1}{2} & \frac{1}{2} \end{pmatrix}$   
 $x_2 = \begin{pmatrix} 3 \\ 3 \end{pmatrix}$   $p_2 = \begin{pmatrix} -\frac{1}{2} & \frac{1}{2} \end{pmatrix}$   
 $x_3 = \begin{pmatrix} 3 \\ 7 \end{pmatrix}$

$$PX = \begin{pmatrix} p_1 \\ \vdots \\ p_m \end{pmatrix} (x_1 \dots x_n) = \begin{pmatrix} p_1 \cdot x_1 & \dots & p_1 \cdot x_n \\ \vdots & \ddots & \vdots \\ p_m \cdot x_1 & \dots & p_m \cdot x_n \end{pmatrix} = (y_1 | \dots | y_n)$$



$$x_1 = \begin{pmatrix} 0 \\ 3 \end{pmatrix}$$

$$p_1 = \begin{pmatrix} \frac{1}{\sqrt{2}} & \frac{1}{\sqrt{2}} \end{pmatrix}$$

$$x_2 = \begin{pmatrix} 3 \\ 3 \end{pmatrix}$$

$$p_2 = \begin{pmatrix} -\frac{1}{\sqrt{2}} & \frac{1}{\sqrt{2}} \end{pmatrix}$$

$$x_3 = \begin{pmatrix} 3 \\ 7 \end{pmatrix}$$

$$y_1 = \begin{pmatrix} p_1 \cdot x_1 \\ p_2 \cdot x_1 \end{pmatrix} = \begin{pmatrix} 1.5 \\ -1.5 \end{pmatrix}$$

$$y_2 = \begin{pmatrix} p_1 \cdot x_2 \\ p_2 \cdot x_2 \end{pmatrix} = \begin{pmatrix} 3 \\ 0 \end{pmatrix}$$

$$y_3 = \begin{pmatrix} p_1 \cdot x_3 \\ p_2 \cdot x_3 \end{pmatrix} = \begin{pmatrix} 5 \\ -2 \end{pmatrix}$$

How to choose P?

- (1) minimise redundancy (measured as covariance)
- (2) maximize the "signal" (measured as variance)

$$C_x \xrightarrow{P} C_Y = \begin{pmatrix} \diagdown & 0 \\ 0 & \diagdown \end{pmatrix} \text{ possible?}$$

Find some (orthonormal) matrix  $P$   
 with  $Y = P \cdot X$ , such that  
 $C_Y = \frac{1}{n} Y Y^T$  is a diagonal matrix.

$$\begin{aligned}
 C_Y &= \frac{1}{n} Y Y^T \\
 &= \frac{1}{n} (PX) (PX)^T \\
 &= \frac{1}{n} P X X^T P^T \\
 &= P \frac{1}{n} X X^T P^T \\
 &= P \cdot C_X \cdot P^T
 \end{aligned}$$

some more linear algebra needed:  
 any symmetric matrix  $A$  can be  
 written as  $A = E \cdot D \cdot E^T$

$(e_1, \dots, e_n)$   $\leftarrow$  orthonormal  
 eigenvectors  $\leftarrow$  diagonal matrix of  
 eigenvalues

choose  $P$ , such that each row  $p_i$  of  $P$  is an  
 eigenvector of  $\frac{1}{n} X X^T = C_X$  (i.e.  $C_X \cdot p_i^T = \lambda_i \cdot p_i^T$ )

$$\begin{aligned}
 C_Y &= P \cdot C_X \cdot P^T \\
 &= P \cdot E \cdot D \cdot E^T \cdot P \\
 &= E^T \cdot E \cdot D \cdot E^T \cdot E
 \end{aligned}$$

even more linear algebra:

1.)  $E$  is orthonormal

2.) the inverse of an orthonormal matrix is the inverse, i.e.  $E^{-1} = E^T$

$$\text{Proof: } (E^T \cdot E)_{ij} = \begin{cases} 1 & \text{if } i=j \\ 0 & \text{else} \end{cases} = I$$

$$\leadsto E^T = E^{-1}$$

$$= E^T \cdot E \cdot D \cdot E^T \cdot E$$

$$= E^{-1} \cdot E \cdot D \cdot E^{-1} \cdot E$$

$$= D$$

cool.