

Section 5.4: Randomized rounding

In Section 5.3 we saw that biasing the prob. of setting each variable true resulted in a better approx. guarantee.

The approximation ratio can be further improved by allowing a different bias for each variable. We will develop an LP-formulation of the problem.

For each clause, C_j , we define:

P_j : the set of indices of variables that occur positively in C_j
 N_j : _____ " _____ negatively _____

Then, C_j can be written as $\bigvee_{i \in P_j} x_i \vee \bigvee_{i \in N_j} \bar{x}_i$

Ex: $\dots \wedge \overbrace{(x_1 \vee \bar{x}_2 \vee \bar{x}_4 \vee x_7)}^{C_j} \wedge \dots$

$$P_j = \{1, 7\}, \quad N_j = \{2, 4\}$$

$$((x_1 \vee x_7) \vee (\bar{x}_2 \vee \bar{x}_4))$$

If $y_i = 0 \Leftrightarrow x_i \equiv F$ and $y_i = 1 \Leftrightarrow x_i \equiv T$,
then C_j is true, iff

$$\sum_{i \in P_j} y_i + \sum_{i \in N_j} (1 - y_i) \geq 1$$

This leads to the following IP-formulation:

IP_ϕ :

$$\max \sum_{j=1}^m z_j w_j$$

subject to

$$\sum_{i \in P_j} y_i + \sum_{i \in N_j} (1 - y_i) \geq z_j, \quad 1 \leq j \leq m$$

$$y_i \in \{0, 1\}, \quad 1 \leq i \leq n$$

$$z_j \in \{0, 1\}, \quad 1 \leq j \leq m$$

Let LP_ϕ be the LP-relaxation of IP_ϕ , i.e.,

LP_ϕ :

$$\max \sum_{j=1}^m z_j w_j$$

subject to

$$\sum_{i \in P_j} y_i + \sum_{i \in N_j} (1 - y_i) \geq z_j, \quad 1 \leq j \leq m$$

$$0 \leq y_i \leq 1, \quad 1 \leq i \leq n$$

$$0 \leq z_j \leq 1, \quad 1 \leq j \leq m$$

Clearly, $\begin{matrix} \nearrow \\ \text{value of opt. sol.} \\ \text{to } LP_\phi \end{matrix} z_{LP_\phi}^* \geq \begin{matrix} \nearrow \\ \text{value of} \\ \text{opt. sol. to} \\ IP_\phi \end{matrix} z_{IP_\phi}^* = \text{OPT} \begin{matrix} \nwarrow \\ \text{value of opt. sol.} \\ \text{to corresponding} \\ \text{MAXSAT problem} \end{matrix}$

RandRounding (ϕ)

$(\vec{y}^*, \vec{z}^*) \leftarrow$ opt. sol. to LP_ϕ

For $i \leftarrow 1$ to n

Set x_i true with prob. y_i^*

$$\left(\text{i.e., } \sum_{LP_\phi}^* = \sum_{j=1}^m z_j^* w_j \right)$$

The approx. ratio of RandRounding is at least $1 - \frac{1}{e} \approx 0.632$.

For proving this, we will use the following two facts:

Fact 5.8 (Arithmetic-geometric mean inequality):

For any $a_1, a_2, \dots, a_k \geq 0$,

$$\left(\prod_{i=1}^k a_i \right)^{1/k} \leq \frac{1}{k} \sum_{i=1}^k a_i$$

\Downarrow

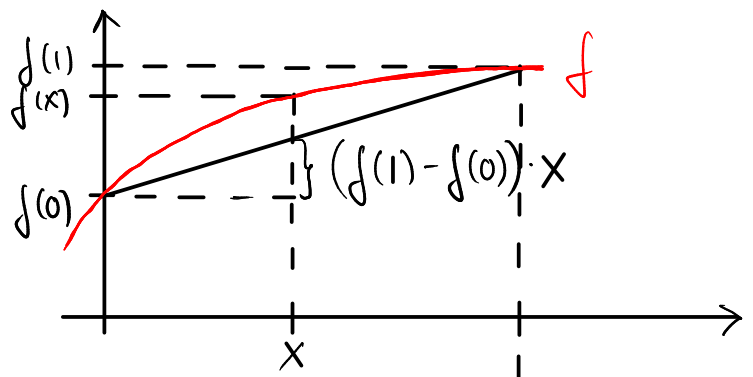
$$\prod_{i=1}^k a_i \leq \left(\frac{1}{k} \sum_{i=1}^k a_i \right)^k$$

A function f is **concave** on an interval I ,
if $f''(x) \leq 0$ for any $x \in I$. (the slope is nonincreasing)

Fact 5.9:

f is concave on $[0, 1]$

$$\Downarrow \forall x \in [0, 1]: f(x) \geq (f(1) - f(0))x + f(0)$$



Theorem 5.10: Round Rounding is a $(1-\frac{1}{e})$ -approx. alg

Proof:

For $1 \leq j \leq m$, let p_j be the probability that C_j is satisfied, and let $\bar{p}_j = 1 - p_j$.

Our goal is to show that $p_j \geq (1 - \frac{1}{e}) z_j^*$.

This will establish the approx factor, since $\text{OPT} \leq \sum_{j=1}^m z_j^* w_j$

$$\begin{aligned}
 \bar{p}_j &= \prod_{i \in P_j} (1 - y_i^*) \prod_{i \in N_j} y_i^* \\
 &\leq \left(\frac{1}{l_j} \left(\sum_{i \in P_j} (1 - y_i^*) + \sum_{i \in N_j} y_i^* \right) \right)^{l_j}, \text{ by Fact 5.8} \\
 &= \left(\frac{1}{l_j} \left(|P_j| - \sum_{i \in P_j} y_i^* + \sum_{i \in N_j} (1 - 1 + y_i^*) \right) \right)^{l_j} \\
 &= \left(\frac{1}{l_j} \left(|P_j| - \sum_{i \in P_j} y_i^* + |N_j| - \sum_{i \in N_j} (1 - y_i^*) \right) \right)^{l_j} \\
 &= \left(1 - \frac{1}{l_j} \left(\sum_{i \in P_j} y_i^* + \sum_{i \in N_j} (1 - y_i^*) \right) \right)^{l_j}, \text{ since } |P_j| + |N_j| = l_j \\
 &\leq \left(1 - \frac{z_j^*}{l_j} \right)^{l_j}, \text{ since } (\vec{y}^*, \vec{z}^*) \text{ is a solution to } LP_\phi
 \end{aligned}$$

Thus,

$$p_j \geq 1 - \left(1 - \frac{z_j^*}{l_j} \right)^{l_j} \equiv f(z_j^*)$$

which is a concave function of z_j^* :

$$f'(z_j^*) = - l_j \left(1 - \frac{z_j^*}{l_j} \right)^{l_j-1} \cdot \left(-\frac{1}{l_j} \right) = \left(1 - \frac{z_j^*}{l_j} \right)^{l_j-1}$$

$$\begin{aligned}
 f''(z_j^*) &= (l_j - 1) \left(1 - \frac{z_j^*}{l_j} \right)^{l_j-2} \cdot \left(-\frac{1}{l_j} \right) = \underbrace{\left(\frac{1}{l_j} - 1 \right)}_{\leq 0} \underbrace{\left(1 - \frac{z_j^*}{l_j} \right)^{l_j-2}}_{\geq 0} \\
 &\leq 0
 \end{aligned}$$

Note that

$$f(0) = 1 - \left(1 - \frac{0}{\lambda_i}\right)^{\lambda_i} = 1 - 1 = 0$$

$$f(1) = 1 - \left(1 - \frac{1}{\lambda_i}\right)^{\lambda_i}$$

Thus,

$$p_j \geq f(z_j^*)$$

$$\geq (f(1) - f(0)) z_j^* + f(0), \text{ by fact 5.9}$$

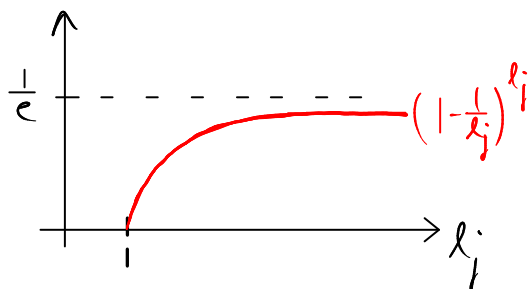
$$\geq \left(1 - \left(1 - \frac{1}{\lambda_i}\right)^{\lambda_i}\right) z_j^*$$

Hence,

$$E[\text{Rand Rounding}] = \sum_{j=1}^m p_j w_j$$

$$\geq \sum_{j=1}^m \left(1 - \left(1 - \frac{1}{\lambda_i}\right)^{\lambda_i}\right) z_j^* w_j$$

$$\geq \left(1 - \frac{1}{e}\right) \cdot \underbrace{\sum_{j=1}^m z_j^* w_j}_{= Z_{LP}^* \geq OPT}$$



□

Note that

Rand Rounding can be derandomized exactly like Rand and Randp

Section 5.5 : Choosing the better of two solutions

Combining the alg.s of Sections 5.1 and 5.4 gives a better approx. factor than using any one of them separately. This is because they have different worst-case inputs:

Rand satisfies clause C_j with prob. $P_R = 1 - (\frac{1}{2})^{l_j}$.

RandRounding satisfies C_j with prob. $P_{RR} \geq (1 - (1 - \frac{1}{l_j})^{l_j}) z_j^*$.

While P_R increases with l_j , the lower bound on P_{RR} decreases with l_j .

BestOfTwo(ϕ)

$\vec{X}_R \leftarrow \text{Rand}(\phi)$

$\vec{X}_{RR} \leftarrow \text{RandRounding}(\phi)$

If $w(\phi, \vec{X}_R) \geq w(\phi, \vec{X}_{RR})$

Return \vec{X}_R

Else

Return \vec{X}_{RR}

Note that

BestOfTwo is **dvandomized** by using the dvandomized versions of Rand and RandRounding.

Theorem 5.11: BestOfTwo is a $\frac{3}{4}$ -approx. alg.

Proof:

$$\begin{aligned} E[\text{BestOfTwo}(\phi)] &= E[\max\{\text{Rand}(\phi), \text{RandRounding}(\phi)\}] \\ &\geq E\left[\frac{1}{2} \text{Rand}(\phi) + \frac{1}{2} \text{RandRounding}(\phi)\right] \\ &= \frac{1}{2} E[\text{Rand}(\phi)] + \frac{1}{2} E[\text{RandRounding}(\phi)], \text{ by lin. of exp.} \\ &\geq \frac{1}{2} \sum_{j=1}^m (1 - 2^{-l_j}) w_j + \frac{1}{2} \sum_{j=1}^m \left(1 - \left(1 - \frac{1}{l_j}\right)^{l_j}\right) z_j^+ w_j \\ &\geq \sum_{j=1}^m z_j^+ w_j \cdot \underbrace{\frac{1}{2} \left(1 - 2^{-l_j} + 1 - \left(1 - \frac{1}{l_j}\right)^{l_j}\right)}_{= p_j}, \text{ since } z_j^+ \leq 1. \end{aligned}$$

$$\text{For } l_j=1, \quad p_j = \frac{1}{2} \left(1 - \frac{1}{2} + 1 - 0\right) = \frac{3}{4}$$

$$\text{For } l_j=2, \quad p_j = \frac{1}{2} \left(1 - \frac{1}{4} + 1 - \left(1 - \frac{1}{2}\right)^2\right) = \frac{3}{4}$$

$$\text{For } l_j \geq 3, \quad p_j \geq \frac{1}{2} \left(1 - \frac{1}{8} + 1 - \frac{1}{e}\right) > \frac{3}{4}$$

Hence,

$$E[\text{BestOfTwo}] \geq \sum_{j=1}^m z_j^+ w_j \cdot \frac{3}{4} \geq \frac{3}{4} \cdot \text{OPT} \quad \square$$